

nature geoscience

NOVEMBER 2012 VOL 5 NO 11
www.nature.com/naturegeoscience

Southern winds and
tree growth trends

EARTHQUAKE TRIGGER
Groundwater extraction

GLACIAL OCEAN
Shallow but strong overturning

WATER ON THE MOON
Traced to the Sun

Recent changes in the Southern Annular Mode are associated with warmer, drier conditions in the Southern Hemisphere. An analysis of tree-ring records there suggests that these changes have significantly altered tree growth. The image shows a millennial-aged cypress tree of the species *Austrocedrus chilensis* in the northern Patagonian Andes

Cover image: Duncan Christie
Cover Design: Karen Moore

Unusual Southern Hemisphere tree growth patterns induced by changes in the Southern Annular Mode

Ricardo Villalba^{1*}, Antonio Lara², Mariano H. Masiokas¹, Rocío Urrutia², Brian H. Luckman³, Gareth J. Marshall⁴, Ignacio A. Mundo¹, Duncan A. Christie², Edward R. Cook⁵, Raphael Neukom^{6,7}, Kathryn Allen⁸, Pavla Fenwick⁹, José A. Boninsegna¹, Ana M. Srur¹, Mariano S. Morales¹, Diego Araneo¹, Jonathan G. Palmer⁹, Emilio Cuq², Juan C. Aravena¹⁰, Andrés Holz¹¹ and Carlos LeQuesne²

Recent changes in the summer climate of the Southern Hemisphere extra-tropics are primarily related to the dominance of the positive phase of the Southern Annular Mode^{1,2}. This shift in the behaviour of the Southern Annular Mode—essentially a measure of the pressure gradient between Southern Hemisphere mid and high latitudes—has been predominantly induced by polar stratospheric ozone depletion^{2–4}. The concomitant southward expansion of the dry subtropical belts^{5,6} could have consequences for forest growth. Here, we use tree-ring records from over 3,000 trees in South America, Tasmania and New Zealand to identify dominant patterns of tree growth in recent centuries. We show that the foremost patterns of growth between 1950 and 2000 differed significantly from those in the previous 250 years. Specifically, growth was higher than the long-term average in the subalpine forests of Tasmania and New Zealand, but lower in the dry-mesic forests of Patagonia. We further demonstrate that variations in the Southern Annular Mode can explain 12–48% of the tree growth anomalies in the latter half of the twentieth century. Tree-ring-based reconstructions of summer Southern Annular Mode indices suggest that the high frequency of the positive phase since the 1950s is unprecedented in the past 600 years. We propose that changes in the Southern Annular Mode have significantly altered tree growth patterns in the Southern Hemisphere.

The most robust and significant change in the Southern Hemisphere atmospheric circulation over recent decades has been the increase in the difference in the mean zonal pressure at sea level between mid and high latitudes^{1,2}. These changes in the pressure fields are well reflected in the increase of the positive phase of the Southern Annular Mode (SAM), also called the Antarctic Oscillation^{1,7}. The persistent positive trend of the SAM in recent decades has been related to a shift to higher latitudes of the Westerlies and associated storm tracks, the poleward migration

of the descending branch of the Hadley cell, and the consequent southern expansion of the dry subtropical belts^{2,5,6}. Several studies have reported significant anomalies in regional temperature and precipitation changes in most Southern Hemisphere continents that are related to changes in the SAM (refs 2,8–11). Tree growth patterns in some specific regions have also been related to inter-annual changes in the SAM (ref. 12). However, no detailed assessment has been carried out of the large-scale impacts of these recent changes in the SAM on the Southern Hemisphere forests or the relative magnitude of these changes in a long-term, multi-century context.

At regional and continental scales, climate is the main driver of forest growth^{13,14}. Tree-ring chronologies provide valuable records of tree growth owing to their annual resolution, wide geographic distribution and temporal coverage over century to millennial timescales¹⁵. The pace of development of tree-ring chronologies has been exponential in the Southern Hemisphere, especially during the past two decades, providing new and updated tree-ring records across South America, Tasmania and New Zealand^{16–18}. Using the largest and most updated network of chronologies available at present for the Southern Hemisphere, we identified regional tree growth patterns for the temperate forests in these three regions during the past 4–600 years (see Supplementary Table S1). The initial analyses of the raw tree-ring width data reveal that the dominant patterns of regional growth observed since the 1950s are significantly different from those recorded in the previous 250 years. Temperate, dry-mesic forests in southern South America show a marked decreasing growth trend over the past five decades that contrasts with the increasing trend recorded in the subalpine trees from Tasmania and New Zealand (Fig. 1).

Regional growth records for each species were developed by merging those individual site chronologies that share a large percentage of common variance. Standardized regional chronologies show that the anomalous tree ring patterns observed

¹Instituto Argentino de Nivología, Glaciología y Ciencias Ambientales, CONICET, CCT-Mendoza, Casilla de Correo 330, 5500 Mendoza, Argentina,

²Laboratorio de Dendrocronología y Cambio Global, Facultad de Ciencias Forestales y Recursos Naturales, Universidad Austral de Chile, Casilla 567, Valdivia, Chile, ³Department of Geography, University of Western Ontario, London, Ontario N6A 3K1, Canada, ⁴British Antarctic Survey, Cambridge, CB3 0ET, UK, ⁵Lamont-Doherty Earth Observatory of Columbia University, Palisades, New York 10964, USA, ⁶Oeschger Centre for Climate Change Research, University of Bern, 3012 Bern, Switzerland, ⁷Swiss Federal Research Institute WSL, 8903 Birmensdorf, Switzerland, ⁸Biological Sciences, Monash University, 310 Blackburn Road, Clayton, Victoria 3800, Australia, ⁹Gondwana Tree-ring Laboratory, PO Box 14, Little River, Canterbury 7546, New Zealand, ¹⁰Centro de Estudios del Cuaternario, 21 de Mayo 1690, Casilla 737, Punta Arenas, Chile, ¹¹School of Plant Science, University of Tasmania, Hobart 7001, Tasmania, Australia. *e-mail: ricardo@mendoza-conicet.gob.ar.

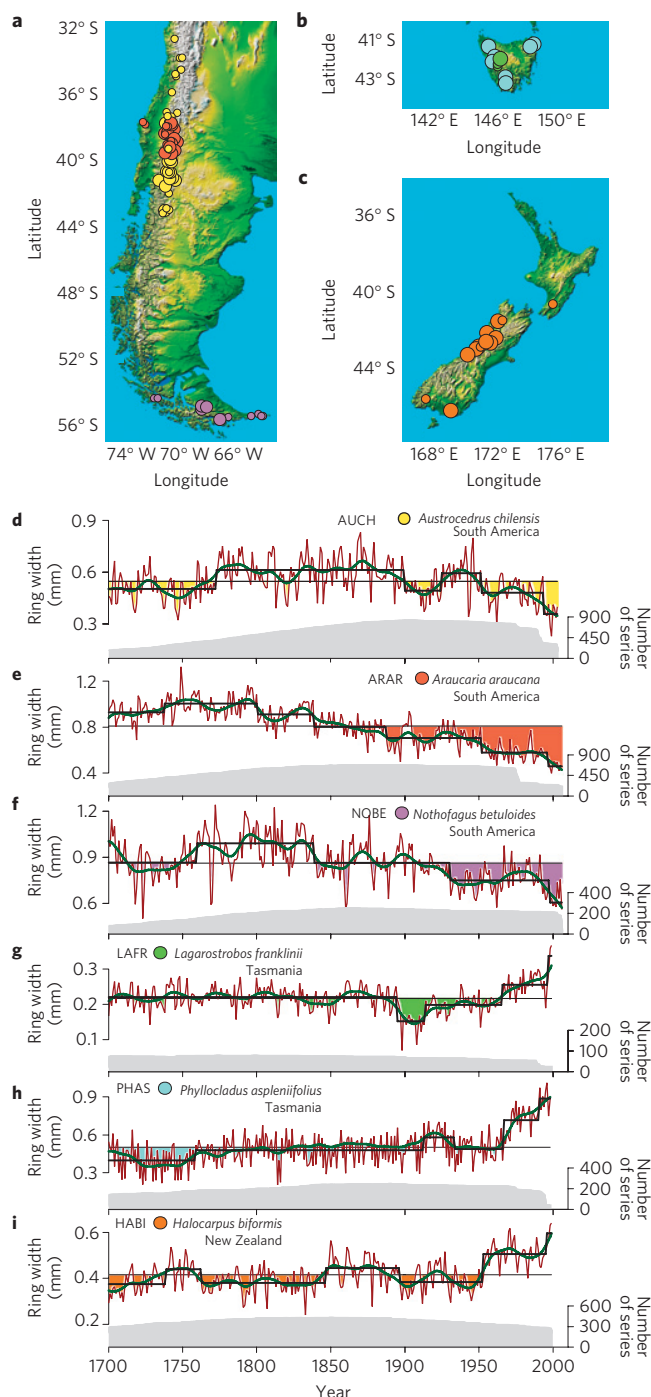


Figure 1 | Long-term variations in tree growth at mid latitudes in the Southern Hemisphere. a–c. Location of tree-ring chronologies.

Chronologies strongly contributing to dominant patterns of tree growth at subcontinental scales are shown with larger circles. **d–i.** Variations in non-standardized regional chronologies since AD 1700. The thick green line emphasizes long-term variations in the raw ring widths. Mean regime levels in tree growth (black lines) were determined using a robust shift detection technique³⁰. The number of tree-ring series in the regional annual means is also shown. Note the recent extreme negative (positive) growth anomalies in Patagonia (Tasmania–New Zealand).

over recent decades were strongly related to regional hydroclimatic records and to the large-scale climatic indices traditionally used to account for variations in the extra-tropical atmospheric circulation in the Southern Hemisphere. Remarkable similarities are observed

between long-term tree growth, climate and the summer SAM index trends during the past 50 years (Fig. 2 and Supplementary Table S3). Recent changes in the radial growth of *Austrocedrus chilensis* and *Araucaria araucana* trees in northwestern Patagonia reflect the gradual decrease in water available at the sampling sites as revealed by significant relationships between tree growth and summer runoff, an integrative estimate of moisture at regional scales (Fig. 2g,h). In southernmost South America, the radial growth of *Nothofagus betuloides* trees has a significant and negative relationship with regional temperatures (that is, suppressed growth during warm years and vice versa), also suggesting that water supply plays a dominant role modulating tree growth at these southern locations (Fig. 2i). The negative correlations of *Austrocedrus*, *Araucaria* and *Nothofagus* tree growth in South America with the SAM index (Fig. 2m–o) are consistent with the decrease in precipitation induced by the SAM in the Patagonian Andes^{9–11}. Growth variations of *Lagarostrobos franklinii* and *Phyllocladus aspleniifolius* trees in Tasmania and *Halocarpus biformis* trees in New Zealand are statistically significantly correlated with temperature changes during the twentieth century in both regions^{16,18} (Fig. 2j–l), which may be in turn partially due to persistent, positive SAM anomalies over the past decades^{11,19} (Fig. 2d–f). In addition, significant relationships have also been recorded between the regional chronologies and the SAM indices reconstructed from the few series of atmospheric pressure available for the mid latitudes in the Southern Hemisphere since the late 1800s (refs 20,21; Fig. 3 and Supplementary Table S3). As shown in Fig. 3a,b,d, trends in tree growth and climate in the Southern Hemisphere mid latitudes are consistent with the climatic anomalies produced by the persistently positive phase of the SAM since the mid 1950s. The significant increase in the atmospheric pressure at mid latitudes of the Southern Hemisphere has been associated with a weakening of the westerly winds, a lower frequency of moisture-laden masses of Pacific origin in South America and a greater number of warm subtropical air masses across New Zealand².

Regression techniques indicate that percentages as high (low) as 48 and 33% (20 and 14%) of the total variance in tree growth variations could be explained by the SAM for the South American and New Zealand forests, respectively (Fig. 2m–o,r). The *Lagarostrobos* and *Phyllocladus* regional records from Tasmania show clear increasing trends during the past decades that we initially related to the positive long-term trend observed in the SAM. However, we noted that the influence of SAM on the radial growth of these species substantially decreases after accounting for autocorrelation in the series, suggesting that a large portion of the significant relationships could be due to the persistent positive trends in both the tree-ring and SAM records (Fig. 2p,q). This suggests that other climatic forcings also play a significant role in modulating the growth of the forests in Tasmania. The weaker relationships observed between records in Tasmania and the SAM are consistent with climatic studies indicating that the warming across Australia during the twentieth century is mainly related to anthropogenic greenhouse gases and sulphate aerosols²². It has been postulated that SAM could have mitigated summer warming in Tasmania^{2,19}. Interestingly, no patterns of significant correlations were recorded between tree growth in the temperate Southern Hemisphere and tropical climatic drivers measured through the Southern Oscillation Index (Supplementary Table S3), suggesting that tree growth in many areas at mid latitudes is largely influenced by extra-tropical circulation.

The tree-ring records discussed in this paper provide a long-term perspective of the common variation in forest growth at mid latitudes in the Southern Hemisphere. The strong relationships observed between these regionally representative records and the SAM series offer the unique opportunity of using the tree-ring

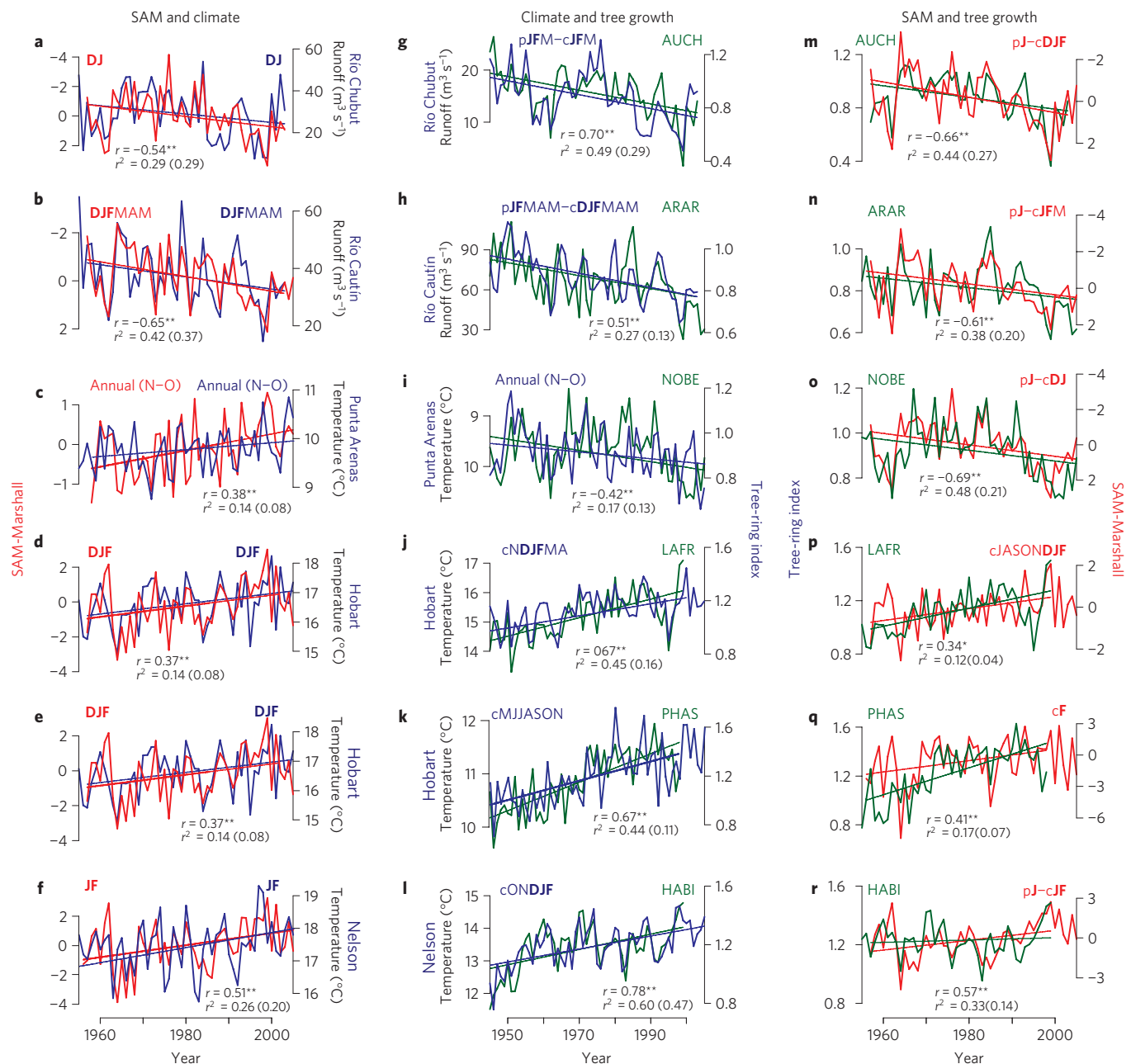


Figure 2 | Influences of the SAM on regional climate and forest growth. a–r. Variables shown include the SAM indices (a–f, m–r; red lines), hydro-climatic variables (a–l; blue lines) and tree-ring indices (g–r; green lines). The left and right columns show correlation coefficients and the percentage of total variance in regional climate and forest growth explained by the SAM. The central column shows similar results for radial growth and regional hydro-meteorological variables. Maximum and minimum (in parentheses) percentages of total variance estimated using the original and the linear-trend-removed series are indicated. Owing to preconditioning of tree growth¹⁵, tree-ring series were compared with climate and the SAM during current (c) and previous (p) growing seasons. SAM indices are plotted inversely in a, b and m–o. Significance at 99% (**) and 95% (*) confidence levels is indicated. Austral summer months (DJF) are highlighted in each seasonal composite.

records as proxy indicators of changes in the SAM during the past centuries. Using time-varying subsets of regional chronologies as predictors, we develop two reconstructions of mean summer (December–February) SAM variations for the period 1409–2006. Two instrumental target series were selected for analysis: the SAM-Marshall⁷ (based on selected station pressure records) and the SAM-NCEP (ref. 1; based on mean sea-level pressure south of 20°S; Fig. 4). The first principal component (PC1) of the three regional chronologies best correlated with the predictands (Fig. 3 and Supplementary Table S3) explains 44 and 47% of the total variance of the SAM-Marshall and SAM-NCEP instrumental

records, respectively (Fig. 4 and Supplementary Table S4). The *Araucaria* regional chronology is the longest record available and alone accounts for 24 and 29% of the total variations in the SAM-Marshall and SAM-NCEP indices, respectively. This regional chronology is used to estimate the SAM variations during the earliest period 1409–1485. For the common period of the three regional chronologies (1559–1999), the mean correlation between the *Araucaria* series and the PC1 based on the *Austrocedrus* and *Halocarpus* chronologies is $r = 0.54$ ($n = 441$, $p < 0.001$), an indication of the high temporal stability between the different predictors used for the reconstruction of the SAM. The temporal

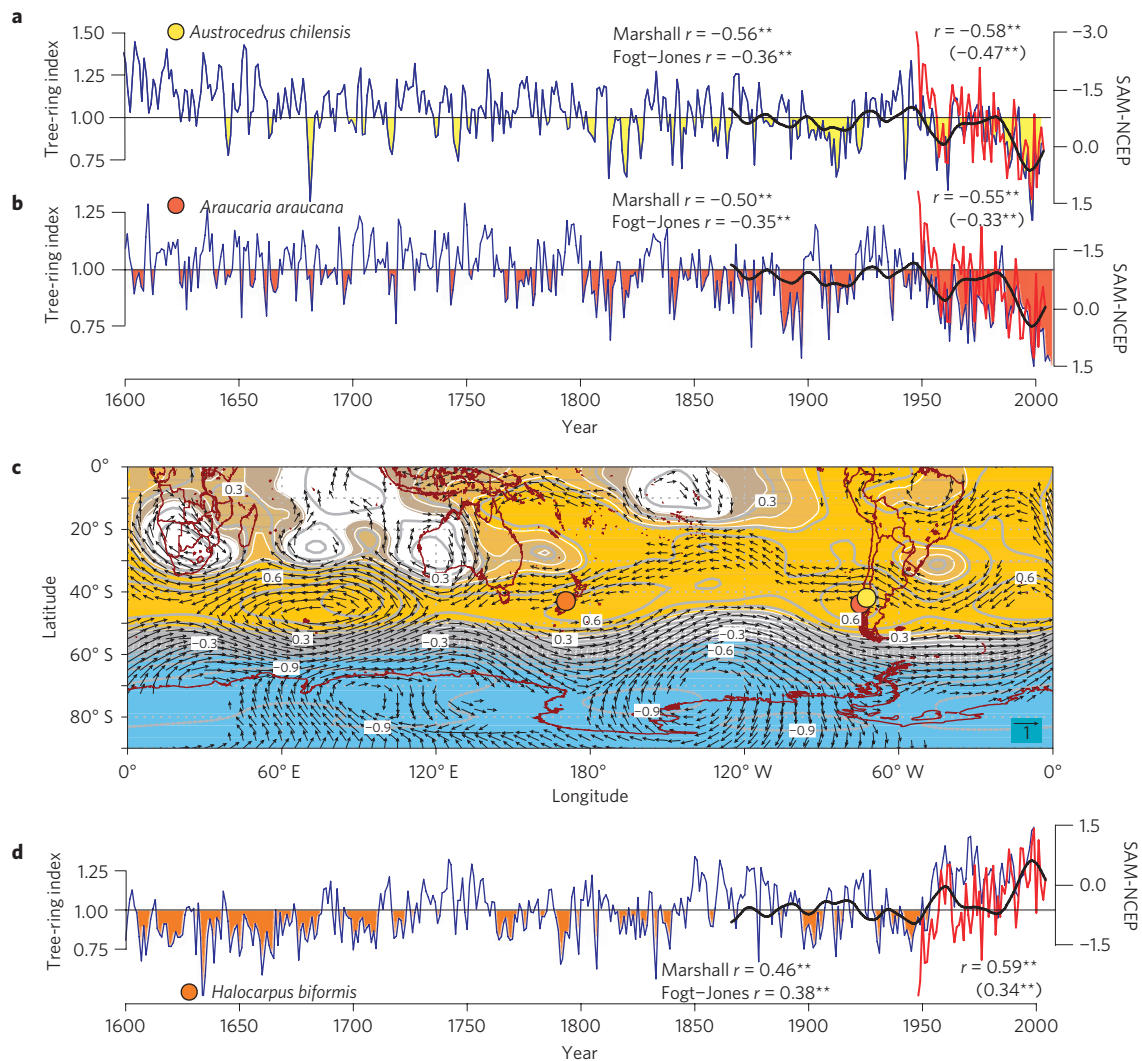


Figure 3 | Variations in regional tree-ring chronologies and the summer SAM. a,b,d, Tree-ring variations since AD 1600 for the three regional chronologies most closely associated with the summer SAM-NCEP (ref. 1; red) and the smoothed mean of the instrumental-based SAM-Fogt²¹/SAM-Jones²⁰ index reconstructions (black). SAM indices are plotted inversely in **a** and **b**. Correlation coefficients over the common period are shown for actual and (trend-removed) series. Coefficients are significant at the 99% confidence level. **c,** Summer correlation fields (1948–2004) between the SAM-NCEP index and mean geopotential heights and wind vectors at 500 hPa. Coloured areas correspond to significant geopotential values at 90 (grey-brown), 95 (brown), 97.5 (brown-orange) and 99% (orange or light blue) confidence levels. Only wind correlation vectors with at least one significant component—**u** (zonal) or **v** (meridional)—at the 95% level are shown. For scale, wind vector (arrow) is at the bottom right. Locations of forests are indicated by circles.

stability and robustness of the SAM reconstructions are also reflected in the high reduction of error (RE) statistics obtained in the regression trials (Fig. 4 and Supplementary Table S4).

The SAM index reconstructions show that after a period of relatively high values during the first half of the fifteenth century, the most extreme negative values of the past 600 years occurred around AD 1470. The SAM indices gradually increased until AD 1500 but recorded negative values until AD 1630–1650. This long period of relatively strong negative values of the SAM index coincides with the most severe phase of the Little Ice Age in southern South America²³. Subsequently, the SAM index values gradually increase and remain relatively stable over the eighteenth century. With the exception of three periods of relatively low index values centred on the decades of AD 1800, 1840 and 1900, the nineteenth and early twentieth centuries show the highest SAM indices for the pre-instrumental period. Relatively low indices of the SAM are recorded from the 1930s to mid 1950s, a period characterized by abundant precipitation in northern Patagonia and lower temperatures in New Zealand^{18,24}. A finer-scale analysis of the most recent decades

indicates that, consistent with the SAM instrumental records, the reconstructions have an initial increase peaking in 1962 with a short decrease centred in 1964–1965. Since this period, both the instrumental and the reconstructed records show a steady increase in trend culminating in 1999–2000 (Fig. 4). Conservative uncertainty boundaries based on a combination of chronology and calibration errors indicate that decadal and longer scale periods (about AD 1450–1680, 1695–1715, 1720–1730, 1785–1800, 1830–1835 and 1930–1940 in the SAM-Marshall reconstruction, and about AD 1425–1800, 1825–1840 and 1925–1950 in the SAM-NCEP reconstruction) significantly differ from the instrumental SAM indices recorded during the most recent decades.

Recent studies based on instrumental records and climatic models reveal that the changes observed in the Southern Hemisphere summer climate during the past five decades have been mainly induced by the ozone depletion in the stratosphere at high latitudes^{2–4}. The development of the ozone hole has not only affected the tropopause over Antarctica but has also induced a poleward shift of the westerly jet at mid latitudes (Fig. 3c), the

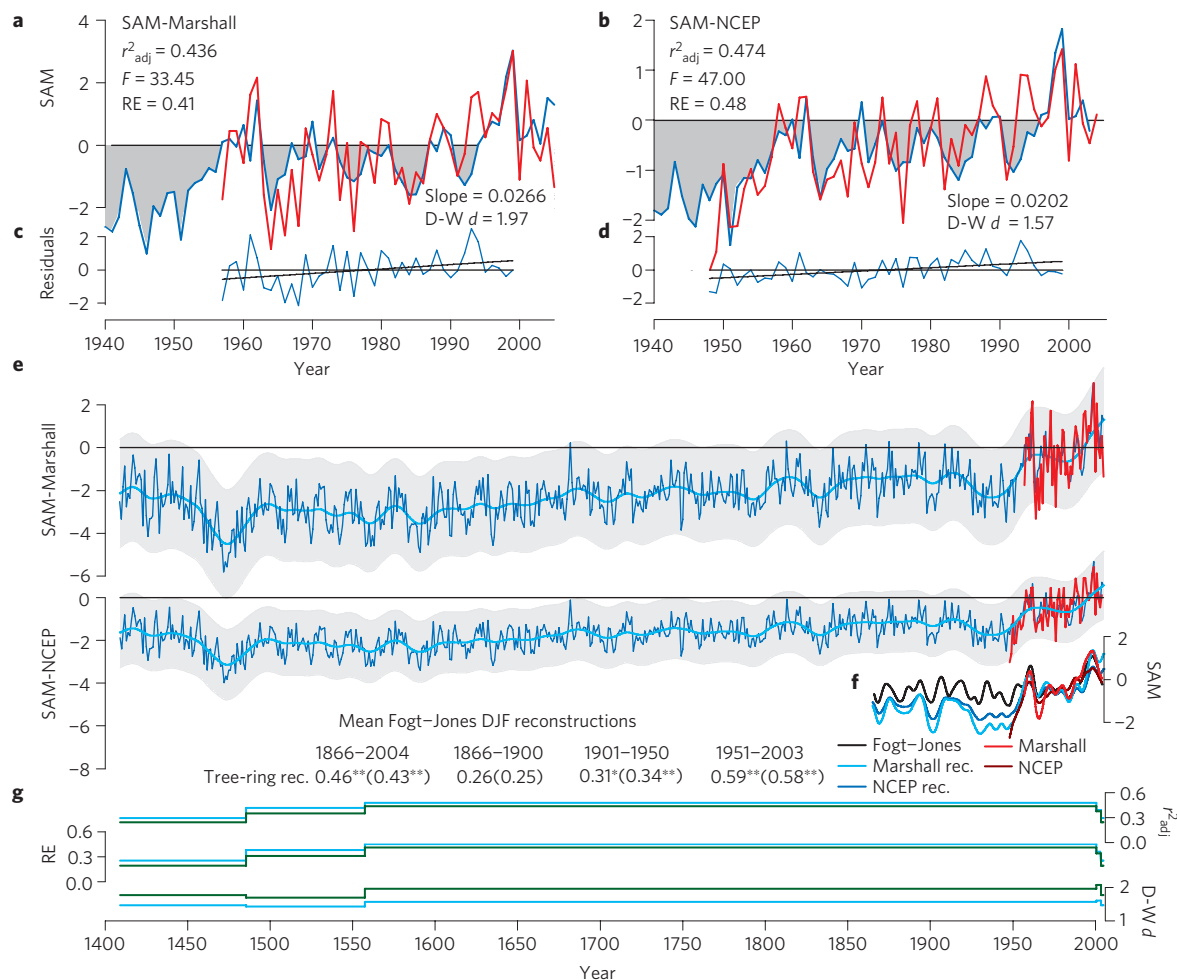


Figure 4 | Summer variations in the SAM for the past 600 years. a–d, Instrumental (red) and tree-ring-based reconstruction (blue) of the summer SAM-Marshall⁷ (**a**) and SAM-NCEP (ref. 1; **b**) with regression residuals (**c,d**). **e**, Nested reconstructions of the summer SAM-Marshall and summer SAM-NCEP for AD 1409–2006 (thick light blue lines, 30-year spline filter) and associated uncertainty bands (grey areas) integrating chronology and calibration errors. **f**, Comparison between the SAM reconstructions, the mean of the instrumental-based reconstructions of the SAM-Fogt/SAM-Jones and the instrumental records. Correlation coefficients between actual and trend-removed (in parenthesis) series and significance (**99%; *95%) are indicated. **g**, Variations in r^2_{adj} , reduction of error (RE) and the Durbin-Watson (DW) d statistic for the SAM-Marshall (green line) and the SAM-NCEP (light blue line) nested models (see Supplementary Table S4 for further information).

displacement of the descending Hadley branch towards higher latitudes, and the poleward migration of the subtropical belts^{2,3,6}. The trends in radial growth of temperate forests since the 1950s reflect the drying and warming patterns observed at mid latitudes in the Southern Hemisphere induced by the recent stratospheric ozone depletion and are consistent with the persistently positive increasing trend in the SAM. Our results show that the recent increase in the SAM indices, as well as the changes in tree growth of temperate forests during past 4–5 decades, seem to be unprecedented in the past 600 years. Further research should address the projected influences of the predicted ozone recovery and the increase in greenhouse gases on the atmospheric circulation in the Southern Hemisphere^{2,3} to elucidate the future patterns of tree growth in Southern Hemisphere temperate forests, which include some of the longest-lived species in the world that are endangered or threatened at present^{16,25}.

Methods

Tree-ring records have been used to determine the influences of temperature, precipitation and other climatic factors on forest growth¹⁵. The database consisted of a recently compiled network of 102 tree-ring sites in the Southern Hemisphere including 39 *Austrocedrus*, 26 *Araucaria*, 13 *Nothofagus*, 14 *Halocarpus*, 7

Phyllocladus and 3 *Lagarostrobos* chronologies (Supplementary Table S1). For each species, site chronologies that share a large percentage of common signal in tree-ring variations were merged in regional chronologies. These regional records are well replicated during the 1980s and 1990s allowing identification of tree growth changes over the past 20 years. There is no tree-ring chronology network from southern Africa of comparable quality and length to those of South America, Tasmania and New Zealand, and therefore no records from that region are included in this analysis.

To conserve the low-frequency signal in the tree-ring records, conservative methods of standardization were selected, fitting negative exponential or linear curves with zero or negative slope to each individual series, or by smoothing each series with a robust median filter²⁶ (for a discussion on the standardization procedures and the different curve fitting see Supplementary Full Methods section and Figs S1–S6).

The SAM-Marshall⁷ (based on a selection of pressure station records from mid and high latitudes of the Southern Hemisphere) and the SAM-NCEP (ref. 1; calculated from sea-level pressure anomalies south of 20° S) index reconstructions were developed following a nested linear regression approach intended to maximize the reconstruction length^{27,28}. The regression statistics of both reconstructions are included in Fig. 4a,b, and in Supplementary Table S4. With the exception of the interval 1409–1485 (77 years) based only on the regional *Araucaria* chronology, all the other reconstructed segments were based on the PC1 of the regional tree-ring records available for each period. The final reconstructions were developed by integrating the segments for each period of the nested models that passed all the verification tests (Supplementary Table S4). This process was conducted after the mean and the variance of each nested reconstruction segment were adjusted to the

best replicated period (1559–1999), which, in turn, registers the highest r^2 value. This procedure was applied to prevent artificial changes in the variance within the reconstruction due to changes in the variance of the predictors²⁸. Two types of error were taken into account to estimate the uncertainties in the SAM reconstructions: the interannual variations in the number and variance of tree-ring series in the regional chronologies (chronology error) and the unexplained variance in the calibration regression model (calibration error)²⁹.

Full methods and associated references are available in the Supplementary Information.

Received 6 June 2012; accepted 21 September 2012;
published online 28 October 2012

References

- Thompson, D. W. J., Wallace, J. M. & Hegerl, G. C. Annular modes in the extratropical circulation. Part II: Trends. *J. Clim.* **13**, 1018–1036 (2000).
- Thompson, D. W. J. *et al.* Signatures of the Antarctic ozone hole in Southern Hemisphere surface climate change. *Nature Geosci.* **4**, 741–749 (2011).
- Polvani, L. M., Waugh, D. J. P., Correa, D. W. & Son, S.-W. Stratospheric ozone depletion: The main driver of twentieth-century atmospheric circulation changes in the Southern Hemisphere. *J. Clim.* **24**, 795–812 (2011).
- McLandress, C. *et al.* Separating the dynamical effects of climate change and ozone depletion. Part II: Southern Hemisphere troposphere. *J. Clim.* **24**, 1850–1868 (2011).
- Archer, C. & Caldeira, K. Historical trends in the jet streams. *Geophys. Res. Lett.* **35**, L08803 (2008).
- Hu, Y. & Fu, Q. Observed poleward expansion of the Hadley circulation since 1979. *Atmos. Chem. Phys.* **7**, 5229–5236 (2007).
- Marshall, G. J. Trends in the southern annular mode from observations and reanalyses. *J. Clim.* **16**, 4134–4143 (2003).
- Gillett, N. P., Kell, T. D. & Jones, P. D. Regional climate impacts of the Southern Annular Mode. *Geophys. Res. Lett.* **33**, L23704 (2006).
- Aravena, J. C. & Luckman, B. H. Spatio-temporal rainfall patterns in southern South America. *Int. J. Climatol.* **29**, 2106–2120 (2009).
- Garreaud, R. D., Vuille, M., Compagnucci, R. & Marengo, J. Present-day South American climate. *Palaeogeogr. Palaeoclimatol. Palaeoecol.* **281**, 180–195 (2009).
- Watterson, I. G. Components of precipitation and temperature anomalies and change associated with modes of the Southern Hemisphere. *Int. J. Climatol.* **29**, 809–826 (2009).
- Lara, A., Villalba, R. & Urrutia, R. A 400-year tree-ring record of the Puelo River summer-fall streamflow in the Valdivian Rainforest eco-region, Chile. *Climatic Change* **86**, 331–356 (2008).
- Boisvenue, C. & Running, S. W. Impacts of climate change on natural forest productivity—evidence since the middle of the 20th century. *Glob. Change Biol.* **12**, 862–882 (2006).
- Davis, M. B. & Shaw, R. G. Range shifts and adaptive responses to Quaternary climate change. *Science* **292**, 673–679 (2001).
- Hughes, M. K., Swetnam, T. W. & Diaz, H. F. *Dendroclimatology, Progress and Prospects. Developments in Paleoenviromental Research 11* (Springer, 2011).
- Cook, E. R. *et al.* Millennia-long tree-ring records from Tasmania and New Zealand: A basis for modelling climate variability and forcing, past, present and future. *J. Quat. Sci.* **21**, 689–699 (2006).
- Boninsegna, J. A. *et al.* Dendroclimatological reconstructions in South America: A review. *Palaeogeogr. Palaeoclimatol. Palaeoecol.* **281**, 210–228 (2009).
- Duncan, R. P., Fenwick, P., Palmer, J. P., McGlone, M. S. & Turney, C. Non-uniform interhemispheric temperature trends over the past 550 years. *Clim. Dynam.* **35**, 1429–1438 (2010).
- Hendon, H. H., Thompson, D. W. J. & Wheeler, M. C. Australian rainfall and surface temperature variations associated with the Southern Hemisphere Annular Mode. *J. Clim.* **20**, 2452–2467 (2007).
- Jones, J. M. *et al.* Historical SAM variability. Part I: Century-length seasonal reconstructions. *J. Clim.* **22**, 5319–5345 (2009).
- Fogt, R. L. *et al.* Historical SAM variability. Part II: Twentieth century variability and trends from reconstructions, observations, and the IPCC AR4 models. *J. Clim.* **22**, 5346–5365 (2009).
- Karoly, D. J. & Braganza, K. Attribution of recent temperature changes in the Australian region. *J. Clim.* **18**, 457–464 (2004).
- Masiokas, M. H. *et al.* Glacier fluctuations in extratropical South America during the past 1000 years. *Palaeogeogr. Palaeoclimatol. Palaeoecol.* **281**, 242–268 (2009).
- Villalba, R. *et al.* Tree-ring based reconstructions of northern Patagonia precipitation since AD 1600. *Holocene* **8**, 659–674 (1998).
- Lara, A. & Villalba, R. A 3,620-year temperature reconstruction from *Fitzroya cupressoides* tree rings in southern South America. *Science* **260**, 1104–1106 (1993).
- Cook, E. R., Briffa, K. R., Meko, D. M., Graybill, D. A. & Funkhouser, G. The segment length curse in long tree-ring chronology development for paleoclimatic studies. *Holocene* **5**, 229–237 (1995).
- Meko, D. M. Dendroclimatic reconstruction with time varying subsets of tree indices. *J. Clim.* **10**, 687–696 (1997).
- Cook, E. R., Woodhouse, C. A., Eakin, C. M., Meko, D. & Stahle, D. W. Long-term aridity changes in the western United States. *Science* **306**, 1015–1018 (2004).
- Eser, J. *et al.* Long-term drought severity variations in Morocco. *Geophys. Res. Lett.* **34**, L17702 (2007).
- Rodionov, S. N. A sequential algorithm for testing climate regime shifts. *Geophys. Res. Lett.* **31**, L09204 (2004).

Acknowledgements

This work was supported by the Inter-American Institute for Global Change Research (IAI) CRN 2047, which is supported by the US National Science Foundation (Grant GEO-0452325), by the Argentinean Council of Research and Technology (CONICET), the Argentinean Agency for Promotion of Science and Technology (PICT02-186) and by FONDECYT Grant Nos 1090479 and 1120965 from the National Research Fund of Chile. G.J.M. is supported by the UK Natural Environment Research Council through the British Antarctic Survey research programme Polar Science for Planet Earth. K.A. is supported by Australian Research Council grant DP120104320 to P. Baker. This work benefited from discussions with P. Sheppard and A. Bahamondez.

Author contributions

R.V., A.L., M.H.M., R.U., B.H.L. and G.J.M. designed the study, interpreted the data and wrote the paper. J.A.B., G.J.M., D.A., R.N. and A.M.S. assisted with climate data, data analysis and figures. R.V., I.M., M.S.M. and A.M.S. provided tree-ring chronologies for Argentina. A.L., D.A.C., R.U., C.L., J.C.A., E.C. and A.H. provided tree-ring records for Chile. E.R.C., P.F., K.A. and J.G.P. provided tree-ring chronologies for New Zealand and Tasmania. All authors discussed the results and commented jointly on the manuscript.

Additional information

Supplementary information is available in the online version of the paper. Reprints and permissions information is available online at www.nature.com/reprints. Correspondence and requests for materials should be addressed to R.V.

Competing financial interests

The authors declare no competing financial interests.

CLIMATE SCIENCE

Tree rings and storm tracks

Reconstructing past climate in the Southern Hemisphere is a challenge. An analysis of tree-ring records suggests that recent changes in the southern storm track in summer are unprecedented in the past 600 years.

Julie Jones

The influence of the Antarctic ozone hole on surface levels of ultraviolet radiation is well known, particularly to sunbathers in the Southern Hemisphere. Less well known is its role in moving the southern storm track towards the South Pole in summer^{1,2}. This large-scale movement of air flow and weather systems has altered surface temperature, precipitation and winds at both mid and high latitudes³. As such, it is important to understand how the recovery of Antarctic ozone levels will influence the storm track, and thus surface climate³. It is expected that the recovery of the ozone hole and increasing greenhouse gas concentrations will exert strong and opposing influences in the coming decades, with ozone recovery shifting the track towards the Equator, and rising greenhouse gas concentrations shifting it towards the South Pole³. Writing in *Nature Geoscience*, Villalba and colleagues⁴ show that changes in the Southern Hemisphere storm track in the latter half of the twentieth century significantly modified tree growth, with consequences for climate reconstruction in the region.

Differences in atmospheric pressure between the mid and high latitudes of the Southern Hemisphere are captured by a mode of climate variability known as the Southern Annular Mode (SAM). In the past four decades or so, atmospheric pressure has declined over the South Pole and risen over the mid-latitudes. As a result, the circumpolar westerly winds and the storm track have strengthened. These changes — collectively known as the positive phase of the SAM — have led, for example, to cooling in East Antarctica, and strong warming in parts of the Antarctic Peninsula^{3,5}. The latter is thought to have contributed to the collapse of the Larsen B ice shelf in 2002⁵.

Records of past changes in the SAM are needed to validate the climate and chemistry-climate models used to project future changes in this mode of climate variability. To date, however, observation-based assessments of SAM behaviour have been restricted to the period of time since 1957, when regular measurements

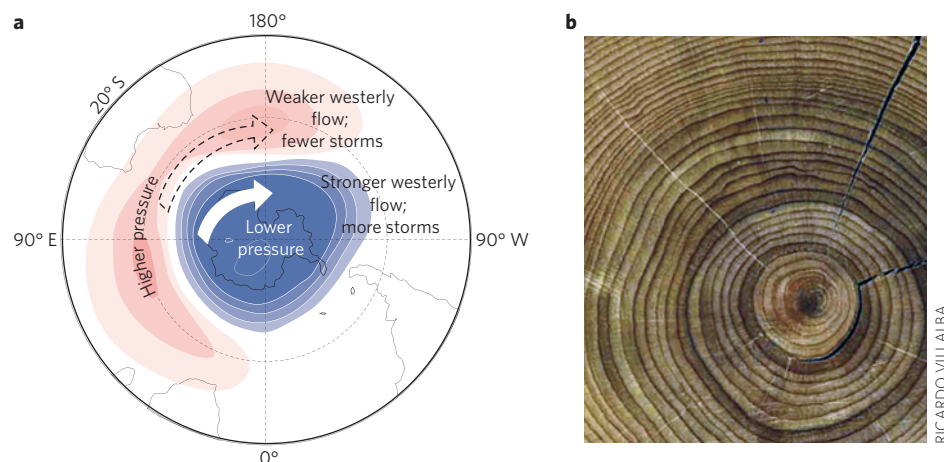


Figure 1 | Positive phase of the Southern Annular Mode. **a**, Differences in atmospheric pressure between the mid and high latitudes of the Southern Hemisphere are captured by a mode of climate variability known as the Southern Annular Mode. The positive phase of this mode, documented here, is characterized by an increase in atmospheric pressure over the mid-latitudes (pink) and a decline in pressure at high latitudes (blue). This pressure difference leads to a strengthening of the circumpolar westerly winds (white arrow) and a southward movement of the Southern Hemisphere storm track. **b**, Villalba and colleagues⁴ use tree-ring records to reconstruct changes in the Southern Annular Mode over the past six centuries, and show that it was more positive at the end of the twentieth century than at any other point over this 600-year period.

in Antarctica began. These records show a trend towards the positive phase of the SAM since the mid-1960s⁵. Extended reconstructions of the SAM are needed to provide a longer-term perspective.

Villalba and colleagues⁴ investigate the influence of ozone-hole-induced changes in the SAM on tree growth over the past few centuries, using tree-ring width data from more than 3,000 trees from Patagonia, New Zealand and Tasmania. They detect a significant shift in tree growth patterns since the 1950s, relative to the previous 250 years, and create regional tree growth records to assess the climatic factors responsible for this shift. According to this approach, the reduction in tree growth in northwestern Patagonia is related to a reduction in summer runoff, and the rise in growth in New Zealand is related to regional warming. These climatic changes, in turn, can be explained by a shift to a more positive phase of the SAM: the higher pressures and

weaker westerly winds in the mid-latitudes that accompany the positive phase of the SAM bring fewer moist Pacific air masses, and thus less precipitation, to northwestern Patagonia, and more warm air masses to much of New Zealand³.

Further exploration of the relationship between tree growth and the SAM, using regression techniques, suggests that variations in SAM behaviour can explain up to 33% and 48% of the variation in tree growth in New Zealand and South America, respectively. On the basis of this relationship, Villalba *et al.* use these tree-ring records to reconstruct changes in the SAM. Their reconstruction extends over the past 600 years, and represents the longest proxy-based record of SAM variability with annual temporal resolution to date. Even when errors are taken into account, the reconstruction suggests that the SAM was more positive at the end of the twentieth century than at any other time over this 600-year period.

This tree-ring-based reconstruction of the SAM supports the results of model simulations and shorter measurement-based records, which also suggest that current values are outside the range of natural variability⁶. Furthermore, Villalba and colleagues demonstrate significant agreement between their reconstructions during the twentieth century and existing ones, giving confidence in their estimates.

Key strengths of their reconstruction of the SAM include the reliance on a large and up-to-date network of trees, the careful analysis of the relationship between tree growth and regional climate, and the production of sound error estimates. The scarcity of tree records extending back to AD 1400 does mean that the earliest parts of the record are based on fewer trees. And as with all reconstructions that rely on relationships between large-scale circulation,

local climate and proxies, the assumption is made that these relationships do not change with time. This is difficult to test — if the long time series were available, we would not need reconstructions in the first place. However, it has been suggested that links between SAM and local climate can change through time in some regions^{7,8}. This is a particular issue for the Antarctic Peninsula region⁷, where the tropical Pacific Ocean modulates the SAM–climate signal, with implications for reconstructions of the SAM reliant on ice-core isotope data from the region.

Villalba and colleagues⁴ utilize tree growth signals to show that changes in the summer SAM in the latter half of the twentieth century are unprecedented in the past 600 years. The vast areas of ocean and the inhospitality of Antarctica make reconstructing climate in the Southern

Hemisphere a challenging yet fascinating task. Climate reconstructions using new proxy data sources, such as those presented in this study, will help to address this challenge⁹. □

Julie Jones is in the Department of Geography, University of Sheffield, Sheffield, S10 2TN, UK.
e-mail: julie.jones@sheffield.ac.uk

References

1. Gillett, N. P. & Thompson, D. W. J. *Science* **302**, 273–275 (2003).
2. Polvani, L. M., Waugh, D. W., Correa, C. J. P. & Son, S. W. *J. Clim.* **24**, 795–812 (2011).
3. Thompson, D. W. J. *et al. Nature Geosci.* **4**, 741–749 (2011).
4. Villalba, R. *et al. Nature Geosci.* <http://dx.doi.org/10.1038/ngeo1613> (2012).
5. Marshall, G. J., Orr, A., van Lipzig, N. P. M. & King, J. C. *J. Clim.* **19**, 5388–5404 (2006).
6. Fogt, R. L. *et al. J. Clim.* **22**, 5346–5365 (2009).
7. Marshall, G. J., Di Battista, S., Naik, S. S. & Thamban, M. *Clim. Dynam.* **36**, 277–287 (2011).
8. Silvestri, G. & Vera, C. *J. Clim.* **22**, 6142–6148 (2009).
9. Neukom, R. & Gergis, J. *Holocene* **22**, 501–524 (2012).

Published online: 28 October 2012

Unusual Southern Hemisphere tree growth patterns induced by changes in the Southern Annular Mode

Full Methods

1. Tree-ring records

We examined interannual variations in tree growth during the last centuries from six species growing at mid-latitude forests in the Southern Hemisphere. The database included an updated and expanded network of tree-ring chronologies from South America, Tasmania and New Zealand. The complete list of chronologies showing their geographic location, the period covered by each record, and its source, is included in Supplementary Information Table 1. These records included only species with adequate sample replication (>20 trees) through the 1980s and 1990s. In South America, the growth patterns were analyzed for three tree species that are representative of temperate forests, the conifers *Austrocedrus chilensis* (AUCH) and *Araucaria araucana* (ARAR) from dry to mesic sites in North Patagonia, and the evergreen broadleaved *Nothofagus betuloides* (NOBE) from relatively mesic to wet sites in southernmost South America. In Tasmania, we assessed the growth patterns of the conifers *Lagarostrobos franklinii* (LAFR) and *Phyllocladus aspleniifolius* (PHAS). Finally, in New Zealand, tree-ring records of *Halocarpus biformis* (HABI) were studied. Except for LAFR, for which we had access to only three chronologies, the records for the other species included a large number of sites distributed, in most cases, across their geographic range of distribution. Therefore, with exception of the LAFR record, the growth variations shown in this study are fairly representative of the regional patterns of tree growth for each selected species (Fig. 1).

In order to identify the dominant patterns of tree growth at a regional scale for each species, tree-ring chronologies were initially standardized using a cubic spline function designed to reduce 50% of the variance in a sine wave with a periodicity of 150 years¹. Principal components (PC) were extracted from the records available for each species over the common period AD 1800–1950, which coincides with the interval of highest sample replication for most chronologies. The selection of this period ensures that the signal in each record is representative of the stand growth, reducing the noise associated with low-series replication in the earlier portion of the chronologies. For each species, all records contributing to PC1 with eigenvalues ≥ 0.60 were combined to develop a regional chronology (Supplementary Information Table 1). This approach was selected as a reasonable compromise between including enough site chronologies to achieve good replication in the regional chronologies while rejecting those chronologies showing site-specific variations due to local fires, logging or insect outbreaks. All the individual series in the regional chronologies were visually

inspected and those exhibiting abrupt changes in growth (releases or suppressions due to stand dynamics) unrelated to climate were removed from the analysis. The use of a relatively lower number of chronologies with a strong internal common signal is the most appropriate method to reduce potential errors associated with the use of larger network of proxy-records that have not been properly assessed for their internal consistency and regional representativeness². Although the three LAFR chronologies had eigenvalues ≥ 0.60 in PC1, only the record from Mount Read was included in the analysis. Previous studies indicated that the records from Lake Marilyn Low and High, located at lower elevations than Mount Read (Supplementary Information Table 1), have different relationships with climate than Mount Read^{3,4}. Although represented only by the Mount Read chronology, the LAFR record includes 176 trees and more than 340 series (see Supplementary Information Table 2). For each of the other five species, the regional chronologies included the following number of site chronologies: AUCH: 20 sites; ARAR: 16 sites; NOBE: 4 sites; PHAS: 7 sites; and HABI: 10 sites. The regional chronologies for each species show a wide range of variation in length, number of series, and tree-ring statistics (see Supplementary Information Table 2).

In order to retain the low-frequency signal in each regional chronology, we tested several standardization methods that utilize different curve fittings^{5,6}. Standardization aims to remove that part of the variability in tree-ring series not related to climate (e.g., tree aging or forest disturbances). Prior to standardization, variations in the standard deviation associated with changes in the mean tree-ring widths over time were stabilized using a data adaptive power transformation⁷. This power transformation has the advantage of homogenizing the variance in the tree-ring series over their entire length, reducing the heteroscedasticity of the raw tree-ring series. After several trials, we selected two standardization methods to preserve the low-frequency variation in the original tree-ring series (Fig. 1). Individual series from raw chronologies showing an increasing trend during the last decades of the 20th century (Fig. 1, LAFR, PHAS, HABI), were standardized using negative exponential or straight lines with zero or negative slope. In contrast, the decreasing trend in the South American raw chronologies (Fig. 1, AUCH, ARAR, NOBE) was better preserved by standardizing each series with a robust median filter, using a time window of 50% of the individual tree-ring series length. Although this standardization method has rarely been used in previous studies, we found that it is the best available option for recovering the decrease in tree growth during the most recent decades. Different curve fittings for each species are shown in Figures 1 to 6 in the Supplementary Information.

These figures show that depending on the standardization method selected, the patterns in tree growth during the most recent decades can be relatively well maintained or severely modified. Since our purpose was to retain low frequency variation in the records during the 20th century, we

use the standardization method that best reproduced the growth trends recorded in the original (raw) data in the most recent decades. For each selected species, trends in tree growth in the original raw data since AD 1900 and the four chronologies using the standardization methods shown in Supplementary Information Figures 1 to 6 are displayed in Full Methods Figure 1. To facilitate comparisons, all records were normalized over the period 1900-2000. Option 2 (negative exponential curve or linear regression with neg. or zero slope) better captures the positive trends in tree growth recorded in LAFR, PHAS and HABI (New Zealand-Tasmania sites) during the 20th century than other standardization options. The regression coefficients for option 2 are closest to the regression coefficients for the raw data in these chronologies (Full Methods Fig. 1). In contrast, option 4 (median smoothing) provides the best alternative to retain the negative trends in tree growth for AUCH, ARAR and NOBE (South American) chronologies during the past few decades. The regression coefficients for option 4 are closest to those recorded in the raw data from these chronologies. Interestingly, none of the selected standardization options spuriously inflate or exacerbate the trends in the raw data in the 20th century. As shown in Full Methods Figure 1, trends in the standardized chronologies are lower (in absolute values) than those in the raw data.

Although increases in chronology mean ages could be invoked as a possible cause for the persistent reduction in tree growth during recent decades in the South American records, the same reasoning is not valid for the Tasmanian-New Zealand chronologies showing similar trends in mean ages as the South American records but opposite patterns in radial growth (Full Methods Fig. 2). As clearly shown in Figure 2, the anomalous patterns in tree growth recorded over most recent decades are consistent with the climate changes recorded at regional scales, which in South America and New Zealand are tightly associated with the persistent trends in the summer SAM since the 1950's. In addition, the mean segment length of the six regional chronologies at the end of the 20th century ranges between 234 and 468 years (Full Methods Fig. 2), suggesting that these segment lengths have little effect on the growth signal on time scales of 2 to 4 decades or even longer.

Multi-decadal shifts in mean conditions in the regional raw chronologies were identified using a simple regime shift detection technique^{8,9} with a cut-off segment length $l = 20$ years, a target probability level $p = 0.05$, and an outlier weighting factor $h = 2$ (Fig. 1). In order to ensure that the identified shifts are not simply the manifestation of a red noise process, each series was prewhitened prior to testing (see references 8 and 9 for further details on this methodology).

2. Climate-tree growth relationships

In order to identify the climate drivers at hemispheric or continental scales associated with the variations in radial tree growth of temperate forests in the Southern Hemisphere, the standardized chro-

nologies were compared with tropical and extra-tropical circulation indices including the Southern Oscillation Index (SOI) and the Southern Annular Mode (SAM) index. The correlation coefficients between the six regional chronologies and the seasonal SAM-Marshall¹⁰ and SAM-NCEP¹¹ indices for the Southern Hemisphere summer (DJF) are shown in Supplementary Information Table 3. In addition, correlation coefficients between the chronologies and the SOI in summer (DJF), autumn (MAM), winter (JJA) and spring (SON) are included. The dominant pattern involves significant correlations between the regional chronologies and the summer SAM index. These are generally negative for the South American species and positive for species from Tasmania and New Zealand. Since the marked trend in the SAM instrumental records (increasing) and the tree-ring chronologies (increasing or decreasing) during the recent decades might inflate the statistical significance of the correlation between records, the long-term trends were removed by using regression techniques for the SAM indices and autoregressive processes for the tree-ring records¹². Correlation coefficients remained significant between the SAM indices and the AUCH, ARAR, NOBE and HABI regional chronologies during summer (DJF, Supplementary Information Table 3) even after accounting for linear trends/autocorrelation in the series, a clear indication of consistent relationships between summer SAM variations and tree-growth. Although significantly correlated with the SAM-Marshall index over the 1957-2007 interval, the NOBE regional chronology has a weaker relationship with the SAM-NCEP index over the 1948-2004 period. The LAFR and PHAS regional chronologies are not significantly correlated with the summer SAM indices once the long-term trends are eliminated (Supplementary Information Table 3).

In addition to testing the strength of the relationships between the SAM and regional tree growth, we also assessed the relationships between (1) the SAM index and the hydro-climate of the regions where the selected species grow, and (2) the regional tree growth and those hydroclimatic parameters affected by the SAM. Our intention here was to show that (1) the SAM influences the hydroclimate in the selected areas and (2) the radial growth of the species of interest is effectively connected with those climate variables affected by the SAM.

In addition, we aim to provide a measure of the percentage of tree growth directly related to the SAM index for each species. Since our previous analyses of the relationships between the SAM index and tree growth (Supplementary Information Table 3) suggested that the strongest relationships occur in summer, the SAM and regional climate conditions were evaluated during the months of December, January and February, or a combination of them. For some species, spring and fall conditions were also included in seasonal means of the SAM, streamflows or temperatures in addition to summer months. As tree growth is commonly influenced by climate during the current and the previous growing seasons¹³, climate and the SAM index during the previous growing season

were also considered in determining the strongest relationships between tree growth, regional climate, and the SAM. For example, the streamflow of the Río Chubut in December-January is clearly modulated by the SAM conditions during these two months ($r = -0.54$; Fig. 2a). The streamflow variations of Río Chubut during the previous and current growing seasons (January to March), an integrative measure of moisture in the dry eastern slopes of the Andes, explain almost 50% of the total variance in AUCH tree growth over the interval 1945-2005 ($r = 0.70$). Finally, the SAM in January of the previous growing season in combination with the SAM index in summer (DJF) of the current growing season explain from 27% up to 44% of the total variations in AUCH tree growth during the common period 1957-2003 (Fig. 2m). A similar analysis indicated that (1) the SAM in summer and fall (December to May) modulated 37 to 42% of the total variations in the Río Cautín streamflows during the past 47 years (Fig. 2b), (2) summer-fall variations of the Río Cautín during previous and current growing seasons are closely related to ARAR tree growth (Fig. 2h), and (3) the SAM in previous January and current January to March explain between 20 and 38% of the total variance in ARAR tree growth since the year 1957 (Fig. 2n). In establishing the range of total variance in radial growth explained by SAM we followed the Marshall et al. (2006) approach¹⁴, using correlations between detrended and non-detrended data to provide estimations of the minimum and maximum percentages in tree growth variations modulated by the SAM, respectively. Consistent with the results shown in Supplementary Information Table 3, our complementary analyses show that the influence of the SAM on regional climates is stronger in northern Patagonia (Fig. 2a and b) and New Zealand (Fig. 2f) than in southern Patagonia (Fig. 2c) and Tasmania (Figs. 2d and e). Unfortunately, the lack of reliable precipitation records hinders the precise determination of the SAM influences in southern Patagonian climate. The low correlation coefficient between the SAM index and summer temperatures in Tasmania is consistent with recent studies suggesting a weaker influence of the SAM on Tasmanian temperature¹⁵. According to these results, the total variation in tree growth explained by the SAM index in summer reached maximum values around or above 40% in the South American forests and 33% in the New Zealand forests. In contrast, the growth of LAFR and PHAS, which are highly correlated with local temperatures in summer-fall and winter-spring, respectively (Figs. 2j and k), were poorly associated with the SAM index (maximum percentages between 12 and 17%; Figs. 2p and q).

All of the regional chronologies, except for PHAS, were significantly correlated with instrumental-based SAM reconstructions starting in 1866^{16,17} (Supplementary Information Table 3). However, consistent with our previous analyses, the regional chronologies from northern Patagonia and New Zealand show the stronger associations with the instrumental-based SAM reconstructions. Given the weaker correlations of PHAS and LRFR chronologies with the instrumental SAM-Marshall and

SAM-NCEP records, and that between NOBE and SAM-NCEP index, these records were not included as predictors of the SAM indices in the tree-ring based reconstructions discussed below. In addition, these four regional chronologies show weaker correlations with the instrumental-based SAM reconstructions than the AUCH, ARAR and HABI chronologies (Supplementary Information Table 3).

3. Tree-ring based SAM reconstructions

Based on our previous analysis, the AUCH, ARAR and HABI chronologies were selected as potential predictors for the SAM index reconstructions. The first principal component (PC1) from the AUCH, ARAR and HABI regional chronologies captured 60% of the total variations in tree growth during the common period 1559-1999 (441 years) and was significantly correlated with the summer SAM instrumental variations. The interval 1599-1999 had good replication with ≥ 30 series in each regional chronology (Fig. 1 and Supplementary Information Table 1). The contribution of each regional chronology to PC1 loadings was higher for the South America (-0.80 and -0.82 for AUCH and ARAR, respectively) than for the New Zealand region (0.69 for HABI), reflecting the highly significant correlation between the two South American records ($r = 0.514$, $n = 441$, $p < 0.001$). The association between this common pattern of tree growth (PC1) in temperate forests of the SH and the summer SAM is highly consistent even using different SAM records. This is valid for the SAM index derived from (a) a selection of pressure station records from mid and high latitudes of the Southern Hemisphere (SAM-Marshall¹⁰, $r = 0.67$, $n = 43$), (b) the mean sea-level pressure south of 20°S (SAM-NCEP¹¹, $r = 0.70$, $n = 52$), as well as from (c) reconstructions based on longer pressure records from stations at mid-latitudes in the Southern Hemisphere (SAM-Fogt¹⁶, $r = 0.45$, $n = 135$, and SAM-Jones¹⁷, $r = 0.44$, $n = 135$). Note that while the early NCEP reanalysis data are poor over high southern latitudes the errors are smallest during summer, although the seasonal trend is 50% higher than in observations¹⁰.

In order to maximize the length of the SAM reconstructions, we used a nested regression approach in which the shortest predictor series (i.e. the regional tree-ring chronology covering the shortest time interval) was progressively excluded from the regression analyses, and the regression procedure repeated again utilizing the remaining, longer regional chronologies^{18,19}. This reconstruction technique enables the use of concatenated PC records to provide longer reconstructions not limited by the shortest record in the set of predictors. Statistics of SAM-Marshall and SAM-NCEP reconstructions are included in Supplementary Information Table 4. In order to develop the nested reconstructions, the PC1 in year t (i.e. without lagging the series) of the following chronologies was used for the intervals indicated in parentheses: AUCH, ARAR, and HABI (1559-1999) and ARAR

and HABI (1486-1558). The SAM reconstruction during the interval 1409-1485 is solely based on the ARAR regional chronology. For the reconstruction of the interval 2000-2003, we used the PC1 from the two South American chronologies (AUCH and ARAR) since the record from New Zealand was not available for this period. The most recent three years (2004-2006) were estimated using the ARAR regional chronology. The regression model based on the PC1 which includes the three regional chronologies explains 47.4 % and 43.6% of the total variance in the SAM-NCEP (1948-1999) and SAM-Marshall (1957-1999) records, respectively. As expected, the percentage of explained variance decreases back in time as the number of chronologies used to generate the PC predictors decreases, reaching 29.1 % and 24.0% of the explained variance for the SAM-NCEP and SAM-Marshall records, respectively, when only the ARAR chronology is used as predictor (Supplementary Information Table 4). All regression models were cross-validated using a leave-one-out approach¹⁸.

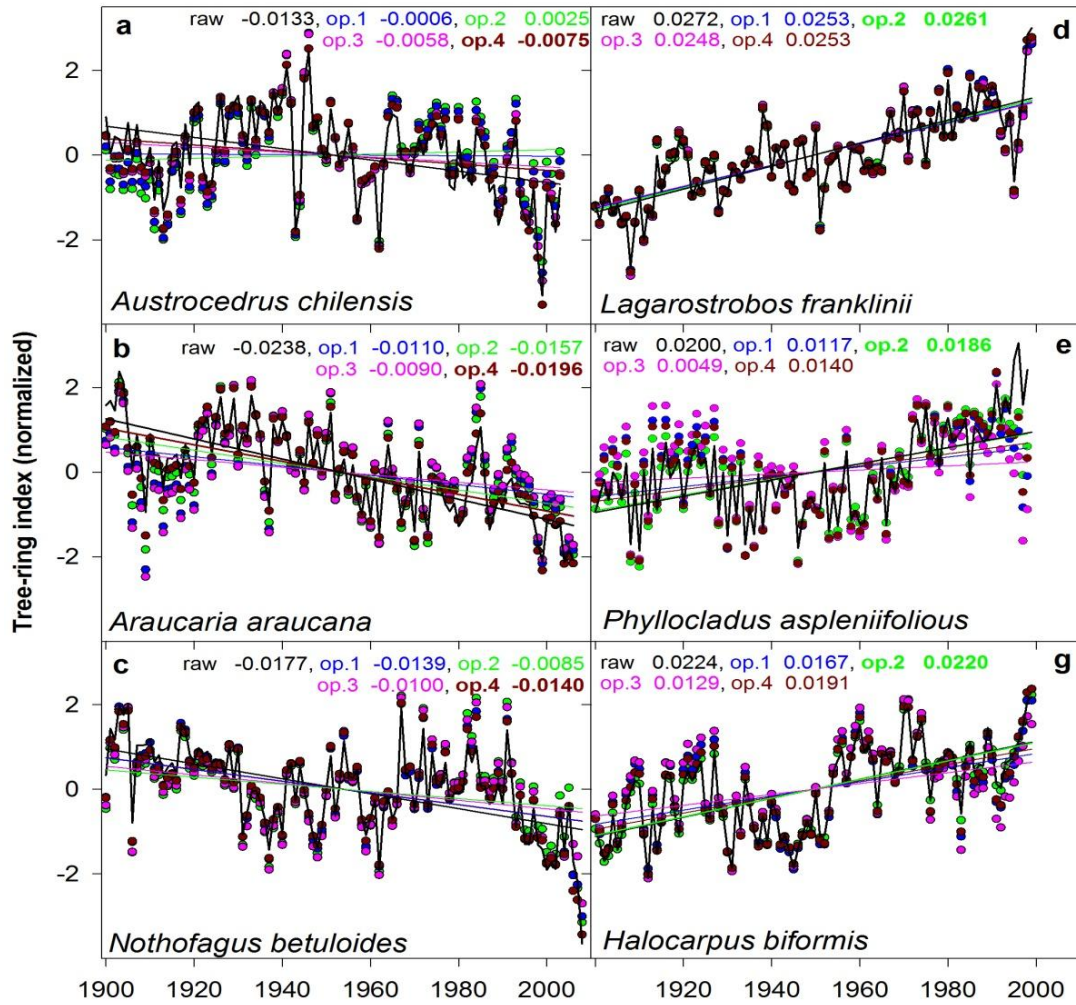
Two types of error were considered in establishing the uncertainties in the SAM reconstructions. These errors included the uncertainties related to changes in the number and variance of regional chronologies back in time (chronology error) and the uncertainty due to unexplained variance in the regression models (calibration error)^{20,21}.

Interannual variations in the standard error of the chronology (S_c), a statistic accounting for both the dispersion around the mean and the varying number of the ring-width measurements per year, was used to quantify the year-to-year variations the chronology error¹³ (Full Methods Fig. 3).

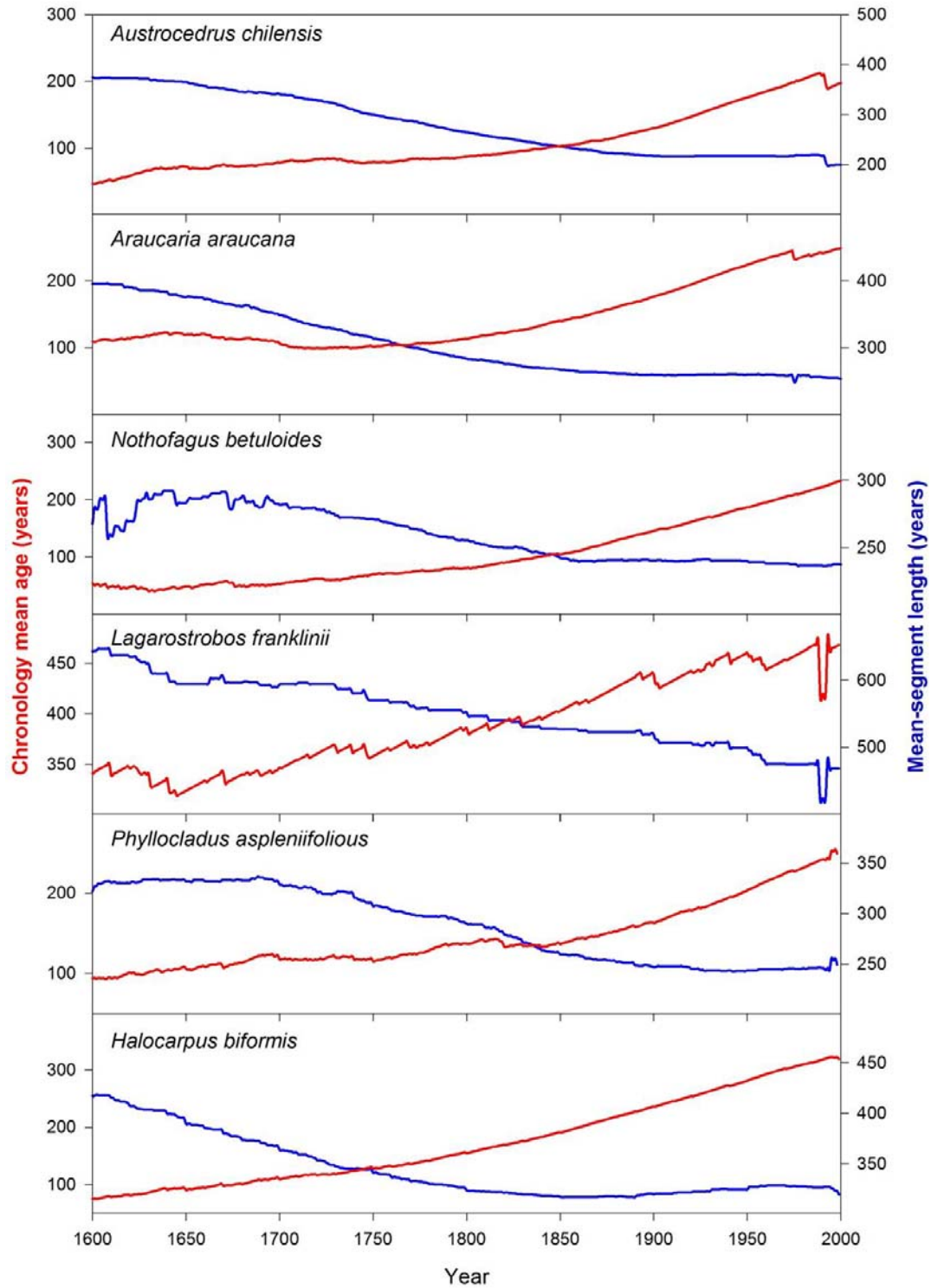
Interannual variations in the S_c of the regional chronologies were transferred to the SAM reconstructions weighting the S_c values by the PC loadings of each chronology in the PC1 used as predictor of the SAM indexes. In the nested SAM index reconstructions, the combined or single S_c associated with the PC1s and the ARAR chronology, respectively, were scaled using the slope regression coefficients from the calibration equations between the PC1 and the SAM instrumental indices.

In addition, three different statistics used to estimate the calibration error in the SAM reconstructions were computed: (1) the standard error of the estimate (S_e), (2) the root-mean-squared error of validation ($RMSE_v$) and the standard deviation of the residuals (SD_{res}) in the calibration models²⁰⁻²². The chronology and calibration errors were quantified separately (Full Methods Fig. 4) and combined (Full Methods Fig. 5) to provide an appreciation of the total errors related to nested SAM-index reconstructions. Due to the extremely high replication in the regional chronologies over the intervals used in the SAM reconstructions (mean replication of 413 samples over the period 1558-2003 for AUCH, 324 samples over the period 1408-2006 for ARAR and 273 samples for HABI over the period 1485-1999), the chronology errors are lower than the calibration errors (Full Methods Fig. 4). The combined uncertainty was estimated from the square root of the summed and

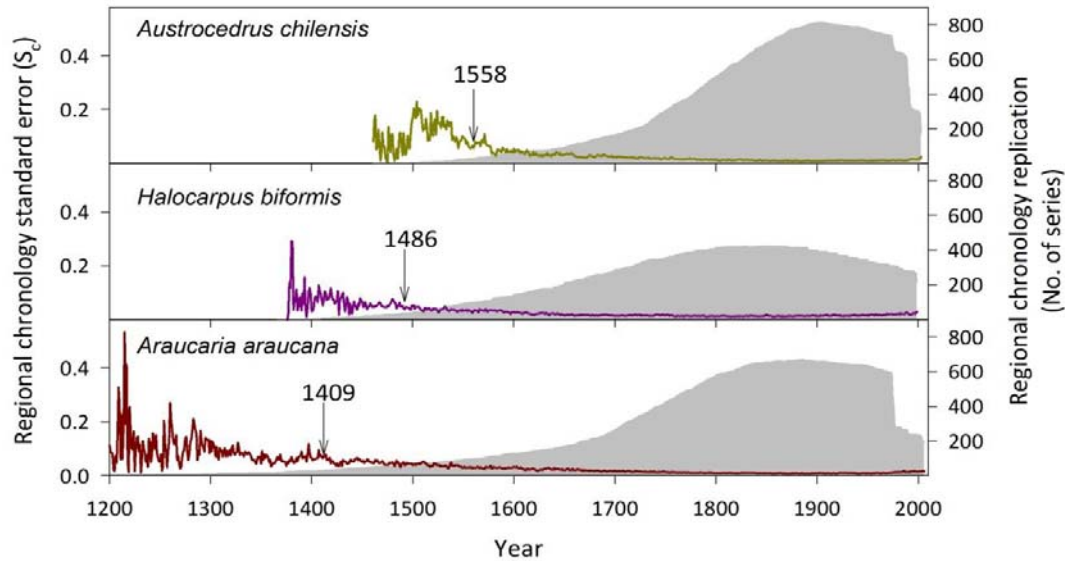
squared chronology and calibration errors²⁰. The integration of the chronology error with the calibration error based on RMSE_v provides the most conservative estimates for the SAM reconstructions (Full Methods Fig. 5). This error combination is displayed in Figure 4 for both the SAM-Marshall and SAM-NCEP index reconstructions.



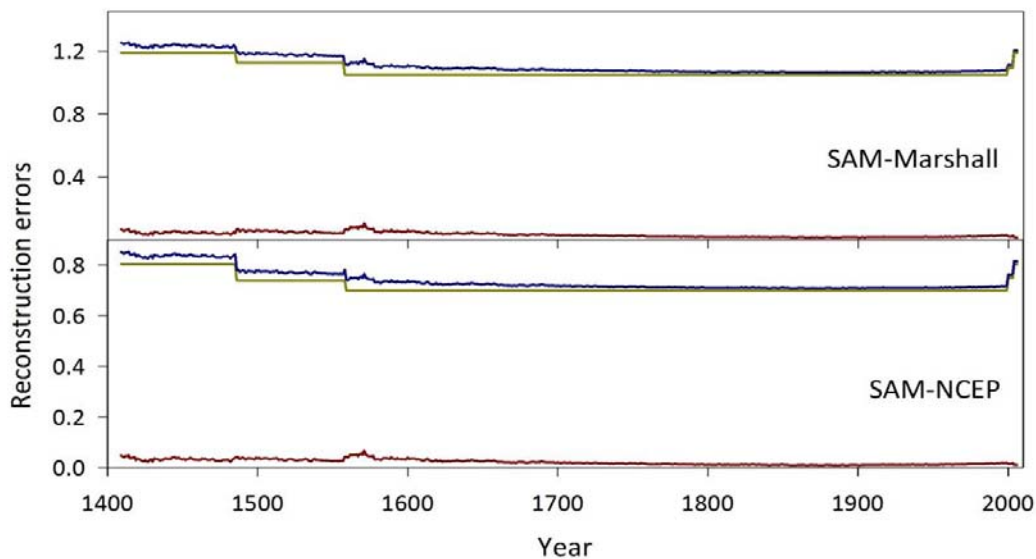
FM_Figure 1. Comparison of normalized tree-growth regression trends during the 20th century between the raw chronologies (solid black line) and the chronologies standardized using four different options. Op. 1 (blue): negative exponential curve or linear regression (any slope); Op. 2 (green): negative exponential curve or linear regression (neg. or zero slope); Op. 3 (pink): general negative exponential curve; Op. 4 (brown): median smoothing. For each species, the regression coefficients are indicated for the raw and the four standardized chronologies. For the South American chronologies (a, b and c), the standardization using median smoothing provides the regression trends closer to those recorded in the raw chronologies. In contrast, the standardization using negative exponential curves or linear regressions with negative or zero slopes show linear trends closer to the raw data for the Tasmania (d and e) and New Zealand (f) chronologies. Note that none of the standardization methods used exacerbates the trends in the raw data.



FM_Figure 2. Patterns of mean age (red) and segment-length (blue) for the tree-ring chronologies used in this study. Note the similarities in both patterns between the different chronologies.

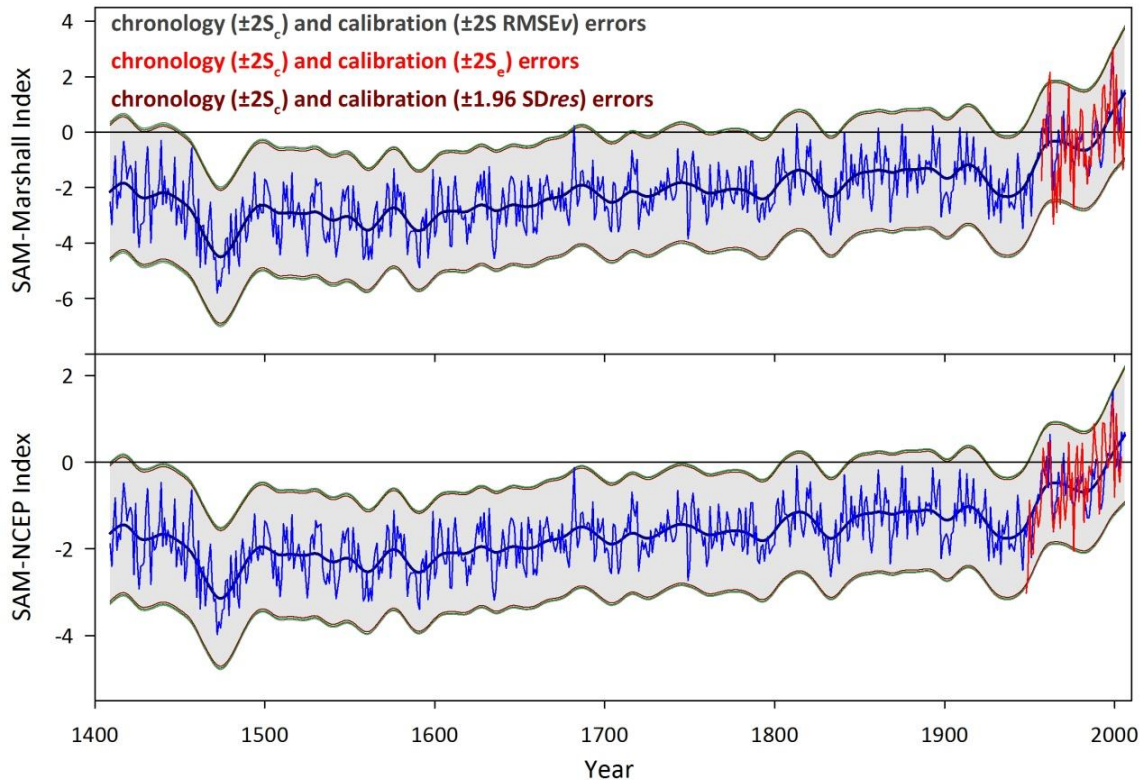


FM_Figure 3. Temporal variations in regional chronology errors (standard error of the chronology, S_c) related to the varying number of ring with measurements back in time and variance of the series around the mean tree-ring index in each chronology. Grey areas represent the replication curves for each regional chronology. Arrows indicate the first year in each chronology with replication ≥ 30 series. We used this criterion to establish the chronology period suitable for developing the reconstructions (more severe than $EPS \geq 0.85$).



FM_Figure 4. Total uncertainties in the SAM-Marshall and SAM-NCEP reconstructions resulting from the sum (blue line) of interannual changes in the number and variance of ring width series around the mean annual ring indices in the regional chronologies (integrated standard error of the chronology S_c ; dark red line) and the unexplained variance in the regression model derived from calibration of the PC1 of the regional chronologies (AUCH, ARAR, HABI over the period 1559-1999; ARAR and HABI over the period 1486-1558; AUCH-ARAR over the years 2000-2003) and the ARAR regional chronology (1409-1487; 2004-2006) against the summer SAM-Marshall and

SAM-NCEP indices (calibration error RMSE_v; dark green line). The integration of the chronology and calibration errors provides an estimation of the overall errors for the nested SAM-index reconstructions.



FM_Figure 5. Summer SAM-Marshal and SAM-NCEP index reconstructions (thin blue lines) together with the 30-year smoothed (thick blue lines) versions back to 1409. Bands around the reconstructions display the errors resulting from the integration of the chronology ($\pm 2S_c$) and calibration uncertainties based on the standard error of estimate ($\pm 2S_c$), the root-mean squared error of validation ($\pm 2RMSE_v$) and the standard deviations of the calibration residuals ($\pm 1.96 SD_{res}$). The three error bands look virtually identical. All error bands smoothed with a 30-year spline filter. Red lines are the instrumental SAM-Marshal and SAM-NCEP indices during austral summer (December to February). Both, the SAM-Marshal and SAM-NCEP indices are dimensionless.

Full Method References

1. Cook, E.R. & Peters, K. The smoothing spline: a new approach standardizing forest interior tree-ring series for dendroclimatic studies. *Tree-Ring Bull.* **41**, 45-53 (1981).
2. Frank, D., Esper, J., Zorita, E. & Wilson, R. A noodle, hockey stick, and spaghetti plate: a perspective on high-resolution paleoclimatology. *WIREs Clim. Change* **1**, 507-516 (2010).

3. Buckley, B.M., Cook, E.R., Peterson, M.J. & Barbetti, M. A changing temperature response with elevation for *Lagarostrobos franklinii* in Tasmania, Australia. *Clim. Change* **36**, 477-498 (1997)
4. Cook, E.R. et al. Millennial-long tree-ring records from Tasmania and New Zealand: A basis for modelling climate variability and forcing, past, present and future. *J. Quaternary Sci.* **21**, 689-699 (2006).
5. Cook, E.R. A time series analysis approach to tree ring standardization. PhD. Dissertation, University of Arizona Tucson, 185 pp. (1985).
6. Briffa, K., & Melvin, T.M. in: *Dendroclimatology, Progress and Prospects. Developments in Paleoenvironmental Research* (eds Hughes, M.K., Swetnam, T.W. & Diaz, H.F.) 113-145 (Springer, Dordrecht, 2011)
7. Cook, E.R. & Peters, K. Calculating unbiased tree-ring indices for the study of climatic and environmental change. *The Holocene* **7**, 361-370 (1997).
8. Rodionov, S.N. A sequential algorithm for testing climate regime shifts. *Geophys. Res. Lett.* **31**, doi: 10.1029/2004GL019448 (2004).
9. Rodionov, S.N. The use of prewhitening in climate regime shift detection. *Geophys. Res. Lett.*, **33**, L12707, doi:10.1029/2006GL025904 (2006).
10. Marshall, G. J. Trends in the southern annular mode from observations and reanalyses. *J. Climate*, **16**, 4134-4143 (2003).
11. Thompson, D. W. J. & Wallace, J. M. Annular modes in the extratropical circulation. Part I: Month-to-month variability. *J. Climate*, **13**, 1000-1016 (2000).
12. Cook, E.R. The decomposition of tree-ring series for environmental studies. *Tree-Ring Bull.* **41**, 45-53 (1987).
13. Fritts, H. *Tree Rings and Climate* (Academic Press, New York, 1976).
14. Marshall, G.J., Orr, A., van Lipzig, N.P.M. & King, J.C. The impact of a changing Southern Hemisphere annular mode on Antarctic Peninsula summer temperatures. *J. Climate*, **19**, 5388-5404 (2006).
15. Hendon, H. H., Thompson, D.W.J. & Wheeler, M.C. Australian rainfall and surface temperature variations associated with the Southern Hemisphere Annular Mode. *J. Climate*, **20**, 2452-2467 (2007).
16. Fogt, R. L. et al. Historical SAM variability. Part II: Twentieth century variability and trends from reconstructions, observations, and the IPCC AR4 models. *J. Climate*, **22**, 5346-5365 (2009).
17. Jones, J. M. et al. Historical SAM variability. Part I: Century-length seasonal reconstructions. *J. Climate* **22**, 5319-5345 (2009).

18. Meko, D.M. Dendroclimatic reconstruction with time varying subsets of tree indices. *J. Climate* **10**, 687-696 (1997).
19. Cook, E.R., Woodhouse, C.A., Eakin, C.M., Meko, D. & Stahle, D.W. Long-term aridity changes in the western United States. *Science* **306**, 1015-1018 (2004).
20. Esper J. et al. Long-term drought severity variations in Morocco. *Geophys. Res. Lett.* **34**, L17702, doi: 10.1029/2007gl030844 (2007).
21. Esper, J. et al. 2012. Orbital forcing of tree-ring data. *Nature Climate Change*, DOI: 10.1038/NCLIMATE1589.
22. Jones, J. M., & Widmann, M. Instrument and tree-ring estimates of the Antarctic Oscillation. *J. Climate* **16**, 3511–3524 (2003).

Supplementary Information

SI_Table 1. Tree-ring chronologies considered in this study. Only chronologies with eigenvalues ≥ 0.6 in the first principal component (PC1 eigenv.) for the period 1800-1950 (shown in bold red), were included in the regional chronology for each species.

| Species/Sites | Code | Latitude/Longitude | Elev. (m) | Period | No. of radii | No. of trees | PC1 eigenv. | Source |
|-------------------------------|-------|--------------------|-----------|-----------|--------------|--------------|-------------|---|
| <i>Austrocedrus chilensis</i> | | | | | | | | |
| El Asiento | ELA | 32°39'S/ 70°49'W | 1800 | 955-2000 | 114 | 94 | 0.09 | LaMarche et al. (1979) ¹ , updated |
| San Gabriel | CH007 | 33°47'S/ 70°15'W | 1600 | 1131-1975 | 63 | 63 | 0.22 | LaMarche et al. (1979) ¹ |
| Río Clarillo | RCL | 33°49'S/ 70°25'W | 1900 | 1204-2002 | 80 | 40 | 0.05 | Le Quesne et al. (2009) ² |
| El Baule | ELB | 34°29'S/ 70°26'W | 1850 | 1146-2001 | 49 | 34 | 0.12 | Le Quesne et al. (2009) ² |
| Agua de la Muerte | AMU | 34°31'S/ 70°25'W | 1850 | 1072-2001 | 108 | 57 | 0.12 | Le Quesne et al. (2009) ² |
| Santa Isabel | CH013 | 34°47'S/ 70°45'W | 850 | 1613-1975 | 64 | 64 | 0.26 | LaMarche et al. (1979) ¹ |
| Alto las Mesas | CH001 | 34°54'S/ 70°42'W | 1020 | 1823-1975 | 7 | 7 | 0.04 | LaMarche et al. (1979) ¹ |
| Melado | CAC | 35°52'S/ 71°00'W | 1435 | 1270-2005 | 108 | 59 | 0.33 | Christie et al. (2011) ³ |
| Polcura | POL | 37°04'S/ 71°24'W | 1205 | 1144-2004 | 154 | 88 | 0.36 | Christie et al. (2011) ³ |
| Huinganco | HUI | 37°07'S/ 70°36'W | 1450 | 190-2000 | 162 | 104 | 0.43 | LaMarche et al. (1979) ¹ , updated |
| El Chacay | CH011 | 37°21'S/ 71°03'W | 925 | 1641-1975 | 16 | 16 | 0.52 | LaMarche et al. (1979) ¹ |
| Abanico | CH010 | 37°21'S/ 71°35'W | 820 | 1733-1975 | 25 | 25 | 0.64 | LaMarche et al. (1979) ¹ |
| Nitrao | NIT | 37°41'S/ 71°17'W | 1220 | 1339-2005 | 119 | 58 | 0.48 | Christie et al. (2011) ³ |
| Ralco-Lepoy | RAL | 38°03'S/ 71°18'W | 1080 | 1585-2005 | 88 | 60 | 0.59 | Christie et al. (2011) ³ |
| Ñorquinco | NOR | 39°07'S/ 71°07'W | 1150 | 1562-2003 | 102 | 52 | 0.69 | Villalba et al. (1998) ⁴ , updated |
| Rucachoroi | RUC | 39°15'S/ 71°10'W | 1300 | 1572-2003 | 50 | 40 | 0.74 | LaMarche et al. (1979) ¹ , updated |
| Quillén | QUI | 39°17'S/ 71°16'W | 1100 | 1676-1989 | 19 | 13 | 0.54 | Villalba et al. (1998) ⁴ |
| Collun-co Alto | COL | 39°56'S/ 71°08'W | 870 | 1596-1989 | 18 | 12 | 0.66 | Villalba et al. (1998) ⁴ |
| Cerro La Hormiga | HOR | 40°03'S/ 71°17'W | 920 | 1508-2003 | 61 | 36 | 0.83 | Villalba et al. (1998) ⁴ , updated |
| Cerro Los Pinos | PIN | 40°04'S/ 71°02'W | 1100 | 1508-2003 | 67 | 44 | 0.62 | Villalba et al. (1998) ⁴ , updated |
| Cerro Castillo | CAS | 40°31'S/ 71°09'W | 1070 | 1576-2003 | 32 | 17 | 0.72 | Villalba et al. (1998) ⁴ , updated |
| El Mirador | MIR | 40°39'S/ 71°24'W | 1050 | 1484-2003 | 52 | 28 | 0.81 | Villalba et al. (1998) ⁴ , updated |
| Piedra del Viento | VIE | 40°42'S/ 71°08'W | 910 | 1679-1991 | 14 | 7 | 0.57 | Villalba et al. (1998) ⁴ |
| Arroyo Minero | MIN | 40°42'S/ 71°16'W | 1050 | 1589-1991 | 16 | 10 | 0.71 | Villalba et al. (1998) ⁴ |
| Confluencia Trafal | CF2 | 40°42'S/ 71°09'W | 1075 | 1723-1989 | 24 | 22 | 0.78 | Villalba et al. (1998) ⁴ |
| Cuyín Manzano | CUY | 40°43'S/ 71°08'W | 900 | 1543-2003 | 39 | 27 | 0.81 | LaMarche et al. (1979) ¹ , updated |
| El Centinela | CEN | 40°44'S/ 71°06'W | 1050 | 1489-2002 | 43 | 27 | 0.85 | Villalba et al. (1998) ⁴ , updated |
| Dedo de Dios | DDD | 40°44'S/ 71°06'W | 900 | 1461-1989 | 37 | 24 | 0.83 | Villalba et al. (1998) ⁴ |
| Cerro El Guanaco | GUA | 41°02'S/ 70°59'W | 1150 | 1497-2002 | 36 | 22 | 0.81 | Villalba et al. (1998) ⁴ , updated |
| San Ramón | RAM | 41°03'S/ 70°59'W | 1100 | 1650-1991 | 17 | 10 | 0.75 | Villalba et al. (1998) ⁴ |
| Cerro Los Leones | LEO | 41°05'S/ 71°09'W | 1020 | 1539-2003 | 73 | 53 | 0.86 | LaMarche et al. (1979) ¹ , Villalba et al. (1998) ⁴ , updated |
| Pilcaniyeu | PIL | 41°11'S/ 70°45'W | 1100 | 1733-1991 | 14 | 7 | 0.72 | Villalba et al. (1998) ⁴ |
| P. del Toro | PAM | 41°32'S/ 71°29'W | 1160 | 1741-1991 | 15 | 15 | 0.65 | Villalba et al. (1998) ⁴ |

| | | | | | | | | |
|------------------------------|------|------------------|------|-----------|-----|----|-------|---|
| La Fragua | FRA | 41°06'S/ 70°54'W | 1180 | 1690-2002 | 23 | 12 | 0.64 | Lab. de Dendrocronología-UACH |
| El Maitén | MAI | 41°59'S/ 71°15'W | 710 | 1690-1974 | 13 | 13 | 0.35 | LaMarche et al. (1979) ¹ |
| Estancia Teresa | TER | 42°57'S/ 71°14'W | 820 | 1540-2002 | 39 | 34 | 0.52 | LaMarche et al. (1979) ¹ , Villalba et al. (1998) ⁴ , updated |
| Nahuel-Pan | NAU | 42°58'S/ 71°13'W | 850 | 1567-2002 | 84 | 49 | 0.58 | Villalba et al. (1998) ⁴ , updated |
| L. Terraplén | RRA | 43°01'S/ 71°34'W | 650 | 1700-1974 | 23 | 23 | 0.62 | LaMarche et al. (1979) ¹ |
| Futaleufú | FUT | 43°11'S/ 71°42'W | 470 | 1700-1992 | 16 | 8 | 0.28 | Villalba et al. (1998) ⁴ |
| <i>Araucaria araucana</i> | | | | | | | | |
| Caramávida | CAR | 37°40'S/ 73°10'W | 900 | 1474-1975 | 21 | 7 | 0.24 | LaMarche et al. (1979) ¹ |
| Copahue | COP | 37°48'S/ 71°04'W | 1720 | 1640-1974 | 10 | 5 | 0.71 | LaMarche et al. (1979) ¹ |
| Piedra del Águila | PIE | 37°49'S/73°01'W | 1300 | 1239-1975 | 35 | 12 | 0.51 | LaMarche et al. (1979) ¹ |
| Caviahue | CAV2 | 37°50'S/ 71°01'W | 1685 | 1444-2003 | 58 | 35 | 0.80 | LaMarche et al. (1979) ¹ , Mundo et al. (2011) ⁵ |
| Ralco | RAL1 | 37°55'S/ 71°22'W | 1280 | 1592-2006 | 26 | 15 | 0.10 | Lab. de Dendrocronología-UACH |
| Vizcacha | VIZ1 | 37°57'S/ 71°17'W | 1440 | 1673-2006 | 24 | 23 | 0.58 | Lab. de Dendrocronología-UACH |
| Chenque Pehuén | CHE | 38°06'S/ 70°51'W | 1616 | 1246-1974 | 42 | 18 | 0.75 | LaMarche et al. (1979) ¹ |
| Nalcas | NAL | 38°19'S/ 71°20'W | 1420 | 1386-1975 | 28 | 8 | 0.66 | LaMarche et al. (1979) ¹ |
| Volcán Lonquimay | LOQ | 38°19'S/ 71°34'W | 1510 | 1664-1975 | 53 | 15 | 0.67 | LaMarche et al. (1979) ¹ |
| Coloradito | COL1 | 38°25'S/ 71°35'W | 1300 | 1612-2006 | 29 | 29 | 0.04 | Lab. de Dendrocronología-UACH |
| Colorado | COLO | 38°25'S/ 71°32'W | 1400 | 1607-2006 | 14 | 10 | 0.01 | Lab. de Dendrocronología-UACH |
| Pino Hachado | HAC2 | 38°40'S/ 70°50'W | 1622 | 1424-2006 | 65 | 36 | 0.80 | LaMarche et al. (1979) ¹ , Mundo et al. (2011) ⁵ |
| Paso del Arco | PAR | 38°49'S/ 71°04'W | 1697 | 1264-2006 | 36 | 19 | 0.40 | Mundo et al. (2011) ⁵ |
| P. Pinos Aluminé | PRI | 38°52'S/ 70°37'W | 1620 | 1140-1974 | 28 | 11 | 0.76 | LaMarche et al. (1979) ¹ |
| Lonco Luan | LON | 38°59'S/ 71°03'W | 1164 | 1306-1974 | 41 | 15 | 0.71 | LaMarche et al. (1979) ¹ |
| Remeco | REM | 39°05'S/ 71°19'W | 1217 | 1450-2006 | 40 | 24 | 0.59 | Mundo et al. (2011) ⁵ |
| Ñorquinco | NOR | 39°09'S/ 71°14'W | 1142 | 1676-2006 | 46 | 27 | 0.65 | Mundo et al. (2011) ⁵ |
| Estancia Pulmarí | PUL | 39°09'S/ 71°13'W | 1181 | 1589-1989 | 28 | 15 | 0.68 | Villalba et al. (1989) ⁶ |
| Lago Rucachoroi | RUC2 | 39°13'S/ 71°09'W | 1317 | 1377-2006 | 130 | 53 | 0.81 | LaMarche et al. (1979) ¹ , Mundo et al. (2011) ⁵ |
| Pinalada Redonda | PINA | 39°19'S/ 71°17'W | 1119 | 1606-2006 | 41 | 25 | 0.57 | Mundo et al. (2011) ⁵ |
| Rahue | RAH2 | 39°23'S/ 70°47'W | 1455 | 1483-2006 | 65 | 35 | 0.76 | LaMarche et al. (1979) ¹ , Mundo et al. (2011) ⁵ |
| Ea. Nahuel Mapi | MAP | 39°32'S/ 71°02'W | 1531 | 1525-2006 | 44 | 25 | 0.86 | Mundo et al. (2011) ⁵ |
| Paso Tromen | TRO2 | 39°36'S/ 71°25'W | 1222 | 1385-1983 | 22 | 12 | 0.75 | Villalba et al. (1989) ⁶ |
| Lago Tromen | TRO | 39°37'S/ 71°20'W | 965 | 1617-1976 | 49 | 18 | 0.68 | LaMarche et al. (1979) ¹ |
| Volcán Lanín | LAN | 39°36'S/ 71°26'W | 1565 | 1291-2006 | 47 | 26 | 0.81 | Mundo et al. (2011) ⁵ |
| Ea Mamuil-Malal | MAM | 39°40'S/ 71°13'W | 890 | 1689-1976 | 26 | 10 | 0.52 | LaMarche et al. (1979) ¹ |
| <i>Nothofagus betuloides</i> | | | | | | | | |
| Beatriz B1 | BEA1 | 53°43'S/ 72°29'W | 45 | 1812-2008 | 31 | 31 | 0.23 | CEQUA |
| Alejandro A0 | ALE0 | 53°44'S/ 72°28'W | 19 | 1704-2008 | 30 | 30 | -0.24 | CEQUA |
| Alejandro A2 | ALE2 | 53°44'S/ 72°28'W | 23 | 1802-2008 | 23 | 23 | -0.39 | CEQUA |
| P. L. Despreciado | LDP | 54°20'S/ 68°49'W | 401 | 1624-2005 | 66 | 55 | 0.84 | Lab. de Dendrocronología-UACH |
| Lago Deseado L.N. | LDN | 54°22'S/ 68°46'W | 455 | 1527-2005 | 52 | 25 | 0.78 | Lab. de Dendrocronología-UACH |
| Lago Fagnano 1 | LFA | 54°28'S/ 68°41'W | 310 | 1577-2005 | 59 | 25 | 0.78 | Lab. de Dendrocronología-UACH |

| | | | | | | | | |
|------------------------------------|-------|-------------------|------|------------|-----|-----|-------------|---------------------------------------|
| Bahía Crossley | AR031 | 54°42'S/ 64°39'W | 20 | 1715-1986 | 15 | 10 | 0.11 | Boninsegna et al. (1989) ⁷ |
| Bahía B. Suceso | AR026 | 54°49'S/ 65°12'W | 35 | 1751-1986 | 15 | 11 | 0.18 | Boninsegna et al. (1989) ⁷ |
| Bahía York | AR024 | 54°49'S/ 64°19'W | 60 | 1647-1986 | 31 | 26 | 0.15 | Boninsegna et al. (1989) ⁷ |
| Puerto Parry | AR051 | 54°49'S/ 64°21'W | 20 | 1726-1986 | 23 | 14 | 0.25 | Boninsegna et al. (1989) ⁷ |
| Río Moat | MOA | 54°53'S/ 66°54'W | 50 | 1528-1986 | 24 | 13 | 0.47 | Boninsegna et al. (1989) ⁷ |
| Ea. Haberton | AR040 | 54°54'S/ 67°19'W | 30 | 1700-1985 | 23 | 17 | 0.22 | Boninsegna et al. (1989) ⁷ |
| Lago Róbalo | LRB | 54°58'S/ 67°41'W | 305 | 1489-2008 | 124 | 76 | 0.80 | Llancabure (2011) ⁸ |
| <i>Lagarostrobos franklinii</i> | | | | | | | | |
| Mt. Read | JMR | 41°51'S/ 145°32'E | 950 | -2145-1999 | 344 | 176 | 0.79 | Cook, et al. (2006) ⁹ |
| Lake Marilyn High | LMH | 42°20'S/ 146°00'E | 800 | 1542-1992 | 23 | 14 | 0.88 | Buckley et al. (1997) ¹⁰ |
| Lake Marilyn Low | LML | 42°12'S/ 146°00'E | 700 | 1058-1992 | 24 | 12 | 0.87 | Buckley et al. (1997) ¹⁰ |
| <i>Phyllocladus aspleniifolius</i> | | | | | | | | |
| Blue Tier | BLT | 41°12'S/ 147°58'E | 750 | 1152-1994 | 84 | 43 | 0.84 | Allen et al. (2001) ¹¹ |
| Ralph's Falls | RFR | 41°19'S/ 147°50'E | 850 | 1240-1994 | 27 | 15 | 0.79 | Allen et al. (2001) ¹¹ |
| Race Course Spur | RCS | 41°29'S/ 145°44'E | 550 | 1467-1994 | 64 | 36 | 0.91 | Allen et al. (2001) ¹¹ |
| Kiaora | KAO | 41°52'S/ 146°06'E | 750 | 1598-1994 | 39 | 29 | 0.78 | Allen et al. (2001) ¹¹ |
| Mt. Arrowsmith | ARR | 42°12'S/ 146°04'E | 800 | 1541-1998 | 36 | 28 | 0.83 | Allen et al. (2001) ¹¹ |
| Scott's Peak Rd. | SPR | 42°50'S/ 146°20'E | 600 | 1519-1995 | 52 | 30 | 0.82 | Allen et al. (2001) ¹¹ |
| Site 600 | SIX | 43°00'S/ 146°33'E | 600 | 1543-1998 | 35 | 23 | 0.83 | Allen (2002) ¹² |
| <i>Halocarpus biformis</i> | | | | | | | | |
| Putara | PUT | 40°42'S/ 175°31'E | 650 | 1646-1993 | 38 | 21 | 0.37 | D'Arrigo et al. (1998) ¹³ |
| Matiri Range | MAT | 41°34'S/ 172°19'E | 1060 | 1508-1999 | 51 | 28 | 0.51 | Duncan et al. (2010) ¹⁴ |
| Mt. Glasgow | GLS | 41°37'S/ 172°02'E | 1050 | 1491-1999 | 48 | 26 | 0.80 | Duncan et al. (2010) ¹⁴ |
| Croesus Track | CRS | 42°17'S/ 171°23'E | 900 | 1486-1999 | 52 | 35 | 0.77 | Duncan et al. (2010) ¹⁴ |
| Mt. Elliot | MEL | 42°30'S/ 171°50'E | 1050 | 1440-1999 | 30 | 16 | 0.90 | Duncan et al. (2010) ¹⁴ |
| Mt. Te Kinga | TKG | 42°39'S/ 171°30'E | 950 | 1450-1999 | 47 | 27 | 0.88 | Duncan et al. (2010) ¹⁴ |
| Mt. French | MTF | 42°40'S/ 171°20'E | 750 | 1367-1999 | 56 | 31 | 0.86 | Duncan et al. (2010) ¹⁴ |
| Camp Creek | CCP | 42°43'S/ 171°34'E | 970 | 1410-1998 | 72 | 53 | 0.87 | Duncan et al. (2010) ¹⁴ |
| Mt. Greenland | MGR | 42°57'S/ 170°49'E | 865 | 1400-1999 | 48 | 26 | 0.87 | Duncan et al. (2010) ¹⁴ |
| Totora Saddle | TOS | 42°59'S/ 170°51'E | 210 | 1590-1998 | 32 | 20 | 0.42 | Duncan et al. (2010) ¹⁴ |
| Mt. Bonar | BON | 43°05'S/ 173°39'E | 850 | 1463-1999 | 51 | 25 | 0.83 | Duncan et al. (2010) ¹⁴ |
| Omoeroa Saddle | OMO | 43°24'S/ 170°06'E | 320 | 1578-1999 | 34 | 20 | 0.63 | Duncan et al. (2010) ¹⁴ |
| Eldrig Peak | ELD | 45°45'S/ 167°28'E | 750 | 1338-1999 | 52 | 26 | 0.46 | Duncan et al. (2010) ¹⁴ |
| Slopedown Hill | SPD | 45°22'S/ 169°03'E | 560 | 1447-1999 | 25 | 12 | 0.70 | Duncan et al. (2010) ¹⁴ |

UACH: Laboratorio de Dendrocronología, Facultad de Ciencias Forestales y Recursos Naturales, Universidad Austral de Chile, Valdivia, Chile.

CEQUA: Centro de Estudios del Cuaternario de Fuego Patagonia y Antártica, Punta Arenas, Chile.

SI_Table 2. Regional tree-ring chronologies.

| <i>Species</i> | <i>Code</i> | <i>No. of series</i> | <i>No. of trees</i> | <i>Period</i> | <i>SD</i> | <i>Rbar</i> | <i>EPS</i> \geq 0.85 | <i>>30 series</i> |
|------------------------------------|-------------|----------------------|---------------------|---------------|-----------|-------------|------------------------|----------------------|
| <i>Austrocedrus chilensis</i> | AUCH | 853 | 554 | 1461-2003 | 0.188 | 0.183 | 1486 | 1558 |
| <i>Araucaria araucana</i> | ARAR | 758 | 386 | 1140-2006 | 0.314 | 0.245 | 1375 | 1409 |
| <i>Nothofagus betuloides</i> | NOBE | 275 | 181 | 1489-2008 | 0.196 | 0.119 | 1615 | 1633 |
| <i>Lagarostrobos franklinii</i> | LAFR | 344 | 176 | -2145-1999 | 0.212 | 0.177 | -1290 | -1096 |
| <i>Phyllocladus aspleniifolius</i> | PHAS | 339 | 204 | 1152-1998 | 0.199 | 0.220 | 1435 | 1458 |
| <i>Halocarpus biformis</i> | HABI | 457 | 271 | 1367-1999 | 0.229 | 0.186 | 1445 | 1486 |

SD: Standard Deviation

Rbar: is the mean correlation coefficient for all possible pairings among tree-ring series from individual cores, computed for a specific common time interval. In this case we used a 50-year window with a 25-year overlap¹⁵.

EPS \geq 0.85: Early year with the Expressed Population Signal statistic \geq 0.85, considered adequate to reflect a reliable common growth signal¹⁶.

> 30 series: Early year with sample replication \geq 30 series. We used this criterion to establish the chronology period suitable for developing the reconstructions (more severe than *EPS* \geq 0.85).

SI_Table 3. Relationships between the regional chronologies and circulation indices from tropical (Southern Oscillation Index, SOI¹⁷) and extra-tropical (SAM) domains. Relationships significant at 95% confidence level are shown in bold red. *raw*: chronology without standardization, *std*: standard chronology, *res*: residual chronology¹⁸.

| Code | SOI (1948-1999) | | | | SAM-Marshall ¹⁹ ; <i>trendless</i> (n) | SAM-NCEP ²⁰ ; <i>trendless</i> (n) | SAM-Fogt ²¹ (n) | SAM-Jones ²² (n) |
|------------------|--------------------|--------------|-------|--------------|--|--|-------------------------------|--------------------------------|
| | DJF | MAM | JJA | SON | DJF | DJF | DJF | DFJ |
| AUCH- <i>raw</i> | 0.09 | 0.09 | -0.05 | -0.05 | -0.57; -0.45 (47) | -0.62; -0.39 (56) | -0.33 (139) | -0.38 (139) |
| AUCH- <i>std</i> | 0.07 | 0.05 | -0.07 | -0.08 | -0.56; -0.46 (47) | -0.58; -0.39 (56) | -0.34 (139) | -0.36 (139) |
| AUCH- <i>res</i> | 0.06 | -0.01 | -0.14 | -0.13 | -0.53; -0.51 (47) | -0.46; -0.47 (56) | -0.28 (139) | -0.29 (139) |
| ARAR- <i>raw</i> | 0.19 | 0.23 | 0.00 | -0.09 | -0.55; -0.41 (50) | -0.61; -0.35 (57) | -0.32 (140) | -0.38 (140) |
| ARAR- <i>std</i> | 0.14 | 0.17 | -0.05 | -0.15 | -0.50; -0.40 (50) | -0.55; -0.37 (57) | -0.34 (140) | -0.35 (140) |
| ARAR- <i>res</i> | 0.19 | 0.25 | 0.02 | -0.07 | -0.45; -0.42 (50) | -0.37; -0.33 (57) | -0.28 (140) | -0.31 (140) |
| NOBE- <i>raw</i> | 0.24 | -0.04 | -0.11 | -0.26 | -0.51; -0.39 (51) | -0.33; -0.19 (57) | -0.19 (140) | -0.27 (140) |
| NOBE- <i>std</i> | 0.23 | -0.07 | -0.13 | -0.26 | -0.50; -0.38 (51) | -0.31; -0.20 (57) | -0.19 (140) | -0.24 (140) |
| NOBE- <i>res</i> | 0.42 | -0.02 | -0.21 | -0.32 | -0.37; -0.37 (51) | -0.30; -0.28 (57) | -0.18 (140) | -0.19 (140) |
| LAFR- <i>raw</i> | -0.17 | -0.19 | -0.05 | 0.04 | 0.34; 0.15 (43) | 0.54; 0.18 (52) | 0.20 (135) | 0.19 (135) |
| LAFR- <i>std</i> | -0.16 | -0.14 | -0.02 | 0.08 | 0.27; 0.11 (43) | 0.51; 0.19 (52) | 0.18 (135) | 0.19 (135) |
| LAFR- <i>res</i> | -0.18 | -0.16 | -0.02 | 0.07 | 0.20; 0.15 (43) | 0.30; 0.20 (52) | 0.18 (135) | 0.18 (135) |
| PHAS- <i>raw</i> | -0.21 | -0.32 | -0.14 | -0.13 | 0.34; 0.08 (42) | 0.45; -0.03 (51) | 0.14 (134) | 0.20 (134) |
| PHAS- <i>std</i> | -0.20 | -0.24 | -0.18 | -0.15 | 0.20; -0.02 (42) | 0.39; -0.04 (51) | 0.08 (134) | 0.15 (134) |
| PHAS- <i>res</i> | -0.10 | -0.13 | -0.14 | -0.12 | 0.06; 0.05 (42) | -0.02; -0.12 (51) | -0.01 (134) | 0.06 (134) |
| HABI- <i>raw</i> | 0.17 | 0.21 | 0.21 | 0.28 | 0.49; 0.48 (43) | 0.63; 0.48 (52) | 0.38 (135) | 0.36 (135) |
| HABI- <i>std</i> | 0.20 | 0.23 | 0.22 | 0.29 | 0.46; 0.46 (43) | 0.59; 0.47 (52) | 0.38 (135) | 0.34 (135) |
| HABI- <i>res</i> | 0.02 | 0.16 | 0.15 | 0.24 | 0.33; 0.33 (43) | 0.29; 0.34 (52) | 0.38 (135) | 0.37 (135) |

SI_Table 4. Southern Annular Mode (SAM) reconstructions. Calibration and verification statistics.

| Calib. period | Predictand | Predictor | Calibration | | | | Verification | | Reconstr. length |
|---------------|----------------------------|-------------------------|-------------|---------------|------|----------------------|--------------|------|------------------|
| | | | Adj r^2 | F (s_e) | DWd | Residual Trend | RMSEv | RE | |
| 1957-1999 | SAM-Marshall ¹⁹ | PC1 of AUCH, ARAR, HABI | 0.436 | 33.45 (1.087) | 1.97 | 0.0266* | 1.1014 | 0.41 | 1559-1999 |
| 1957-2003 | SAM-Marshall | PC1 of ARAR, AUCH | 0.373 | 28.40 (1.131) | 2.08 | 0.0156 ^{ns} | 1.1487 | 0.34 | 1559-2003 |
| 1957-1999 | SAM-Marshall | PC1 of ARAR, HABI | 0.350 | 23.63 (1.167) | 1.70 | 0.0370* | 1.1912 | 0.31 | 1486-1999 |
| 1957-2006 | SAM-Marshall | ARAR | 0.240 | 16.41 (1.230) | 1.78 | 0.0160 ^{ns} | 1.2604 | 0.19 | 1409-2006 |
| 1948-1999 | SAM-NCEP ²⁰ | PC1 of AUCH, ARAR, HABI | 0.474 | 47.00 (0.699) | 1.57 | 0.0202* | 0.7093 | 0.45 | 1559-1999 |
| 1948-2003 | SAM-NCEP | PC1 of ARAR, AUCH | 0.388 | 35.87 (.749) | 1.61 | 0.0184* | 0.7596 | 0.36 | 1559-2003 |
| 1948-1999 | SAM-NCEP | PC1 of ARAR, HABI | 0.413 | 36.89 (0.739) | 1.43 | 0.0230* | 0.7534 | 0.38 | 1486-1999 |
| 1948-2006 | SAM-NCEP | ARAR | 0.291 | 23.75 (0.804) | 1.47 | 0.0214* | 0.8198 | 0.25 | 1409-2006 |

Adj r^2 : Coefficient of determination adjusted for loss of degrees of freedom.

F (s_e): F-value of regression; (s_e) standard error of the estimate

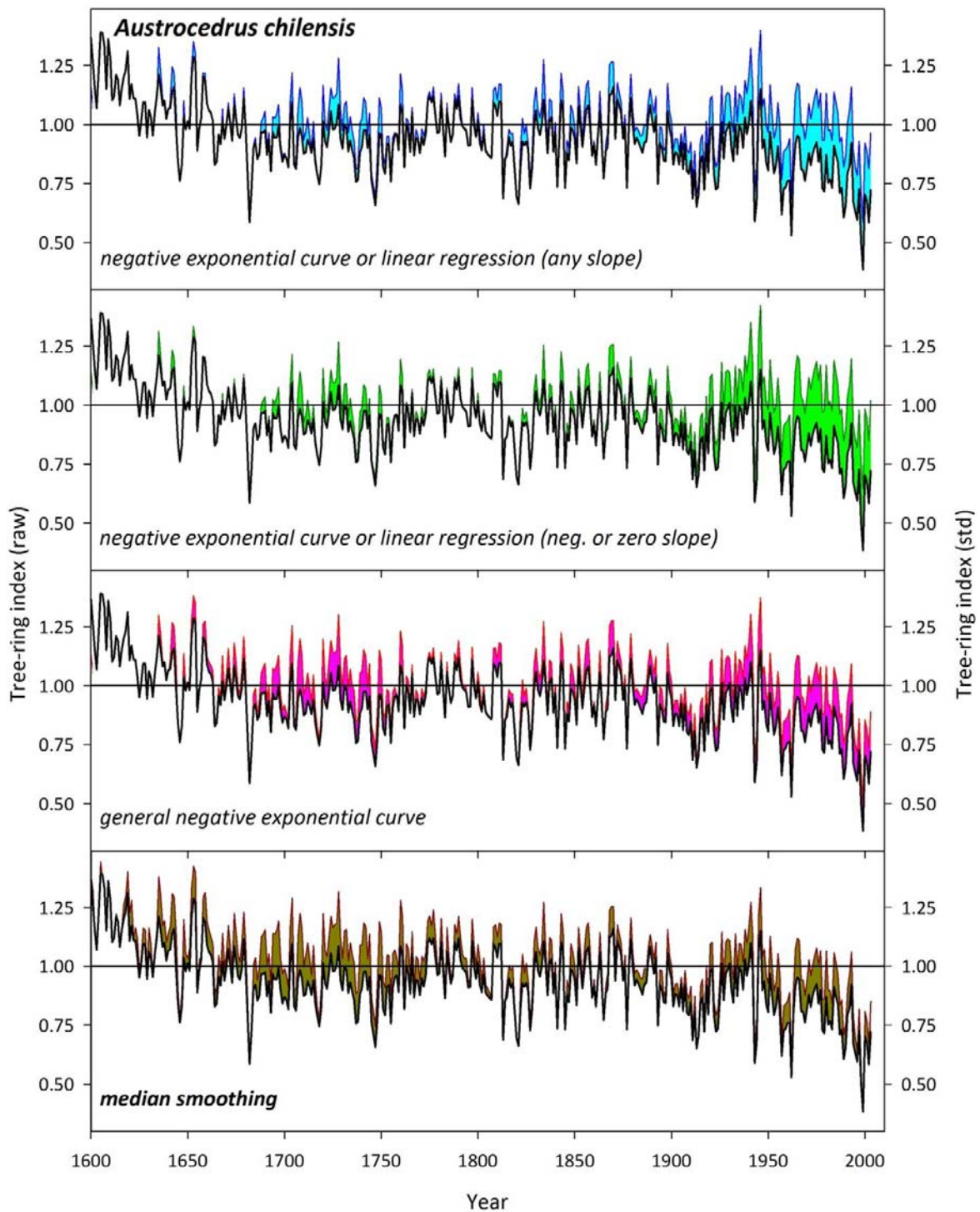
DWd: Durbin-Watson d statistic

RMSEv : the root-mean-square error of cross-validation.

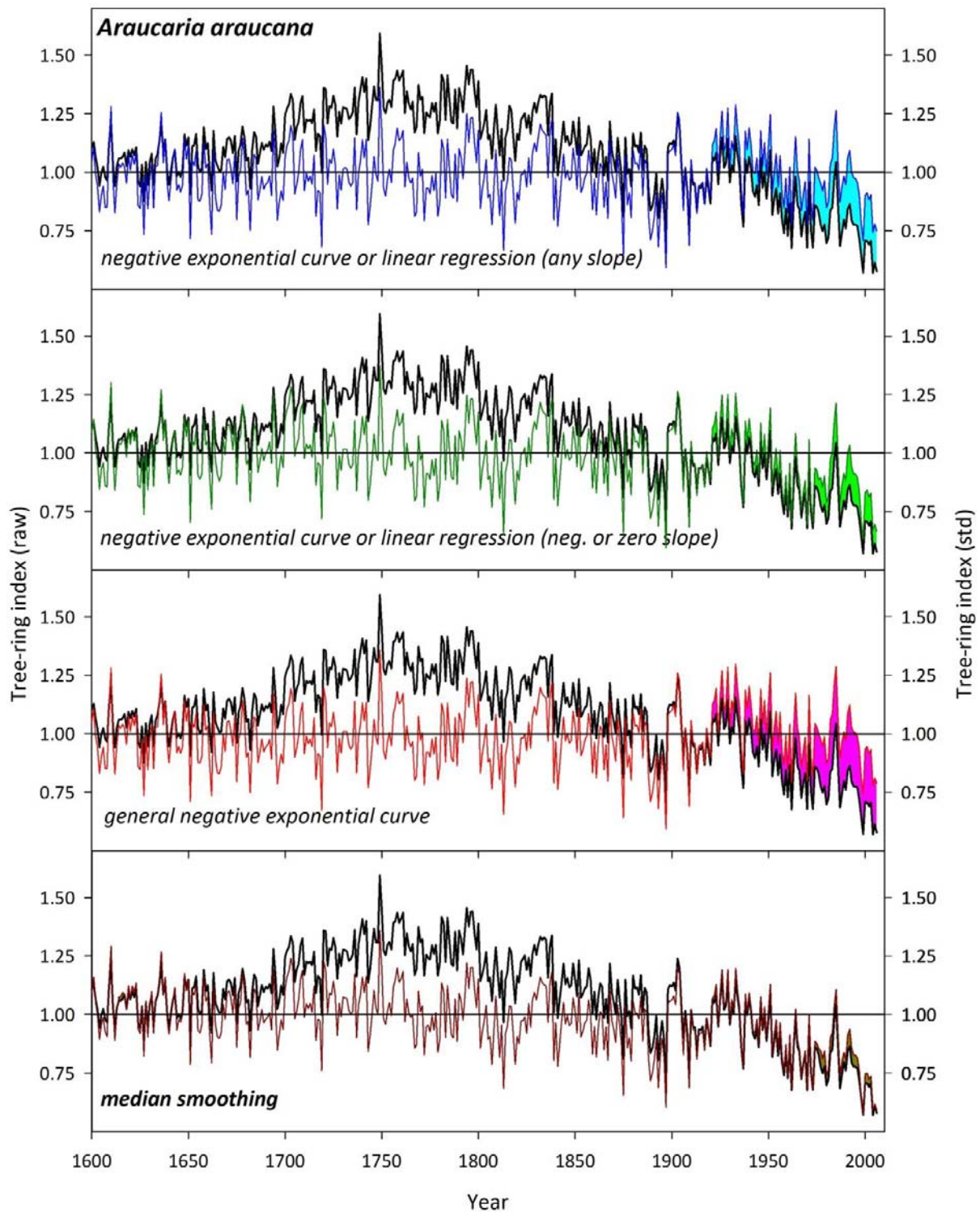
RE: Reduction of Error

*: significant at 95% confidence level.

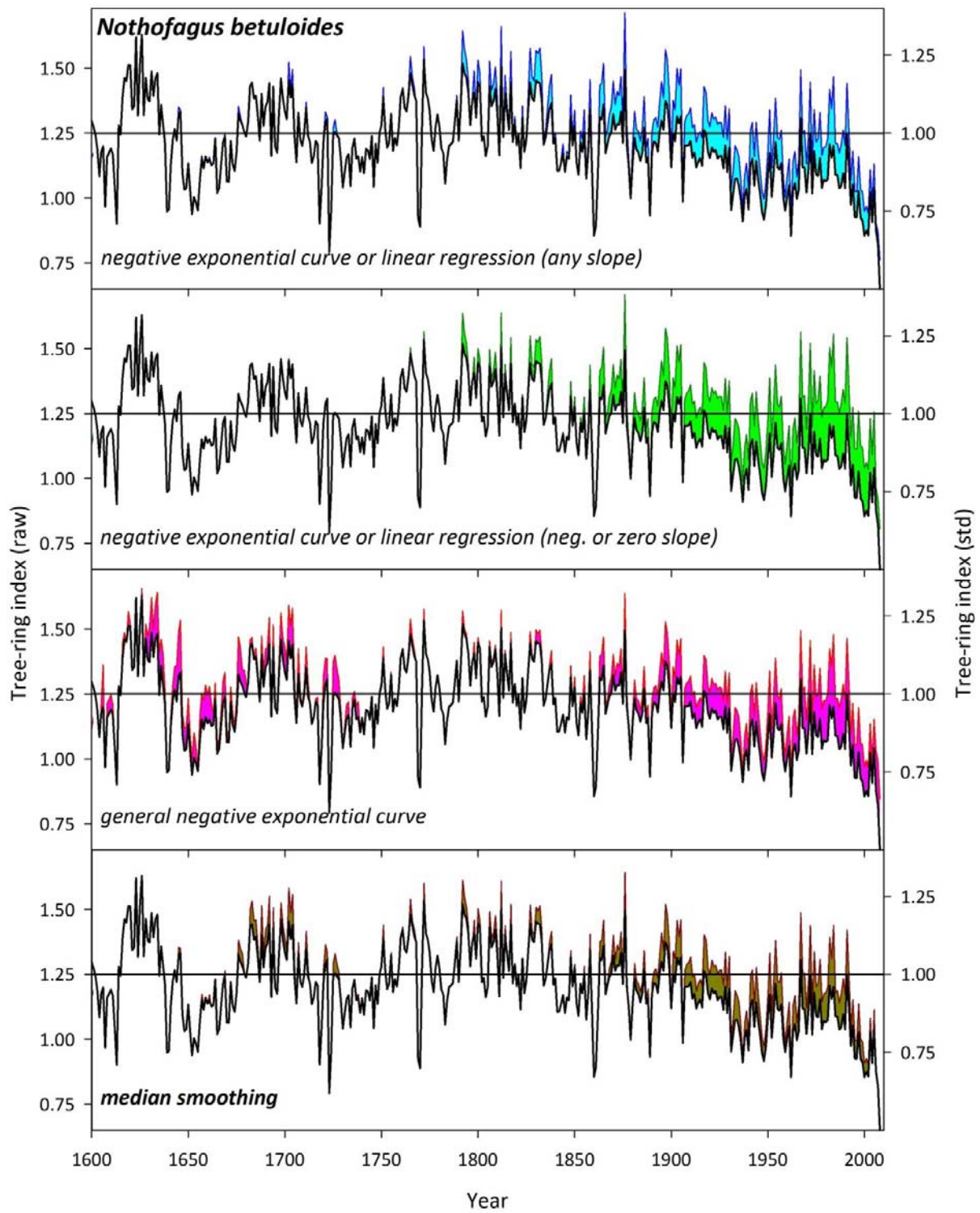
ns: not significant



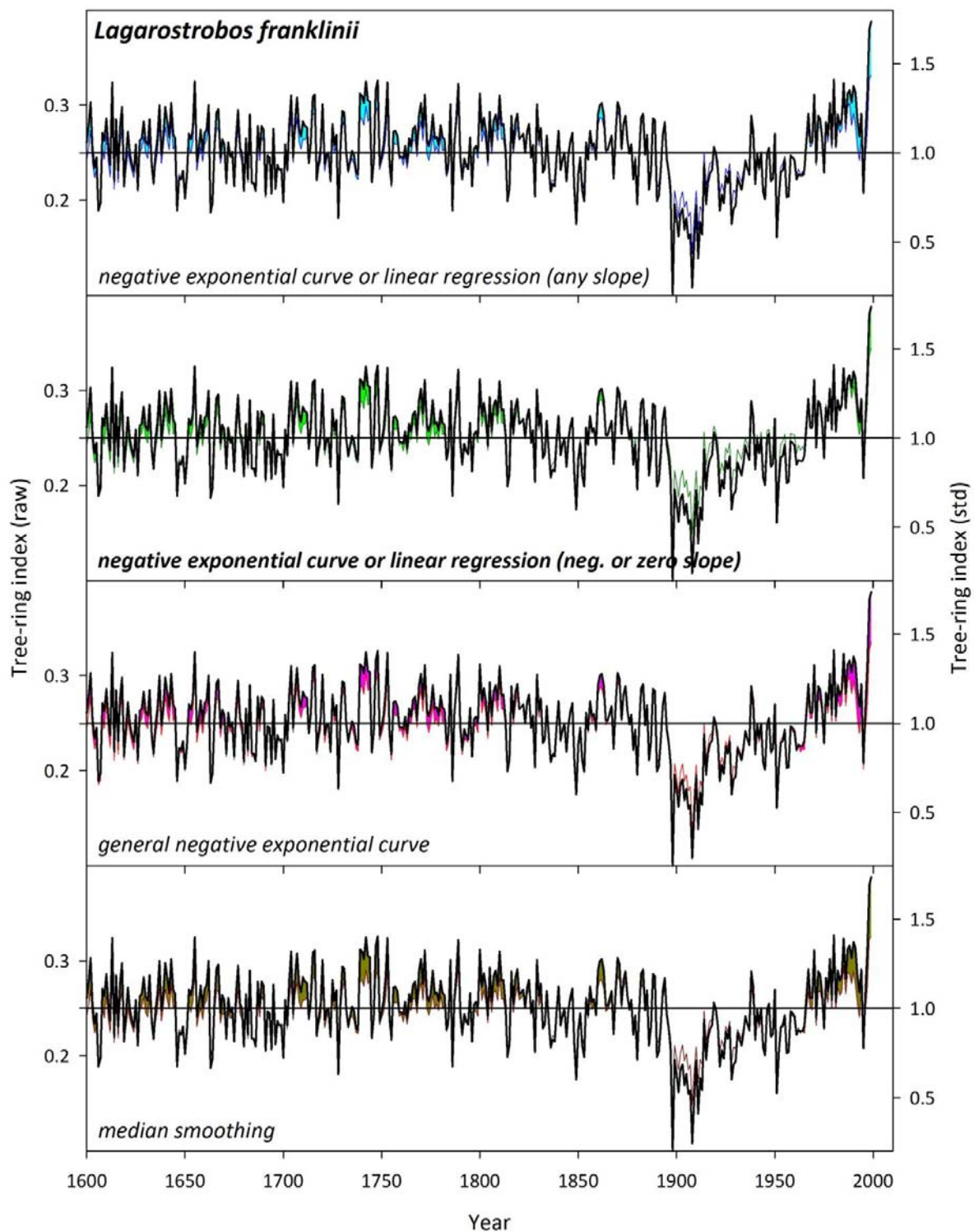
SI_Fig. 1. Comparison between the *Austrocedrus chilensis* raw tree-ring chronology without standardization (black line) and the chronologies produced using different curve fitting methods. The median smoothing method (**highlighted**) captures most of the variance in the low-signal frequency since 1950 contained in the raw chronology.



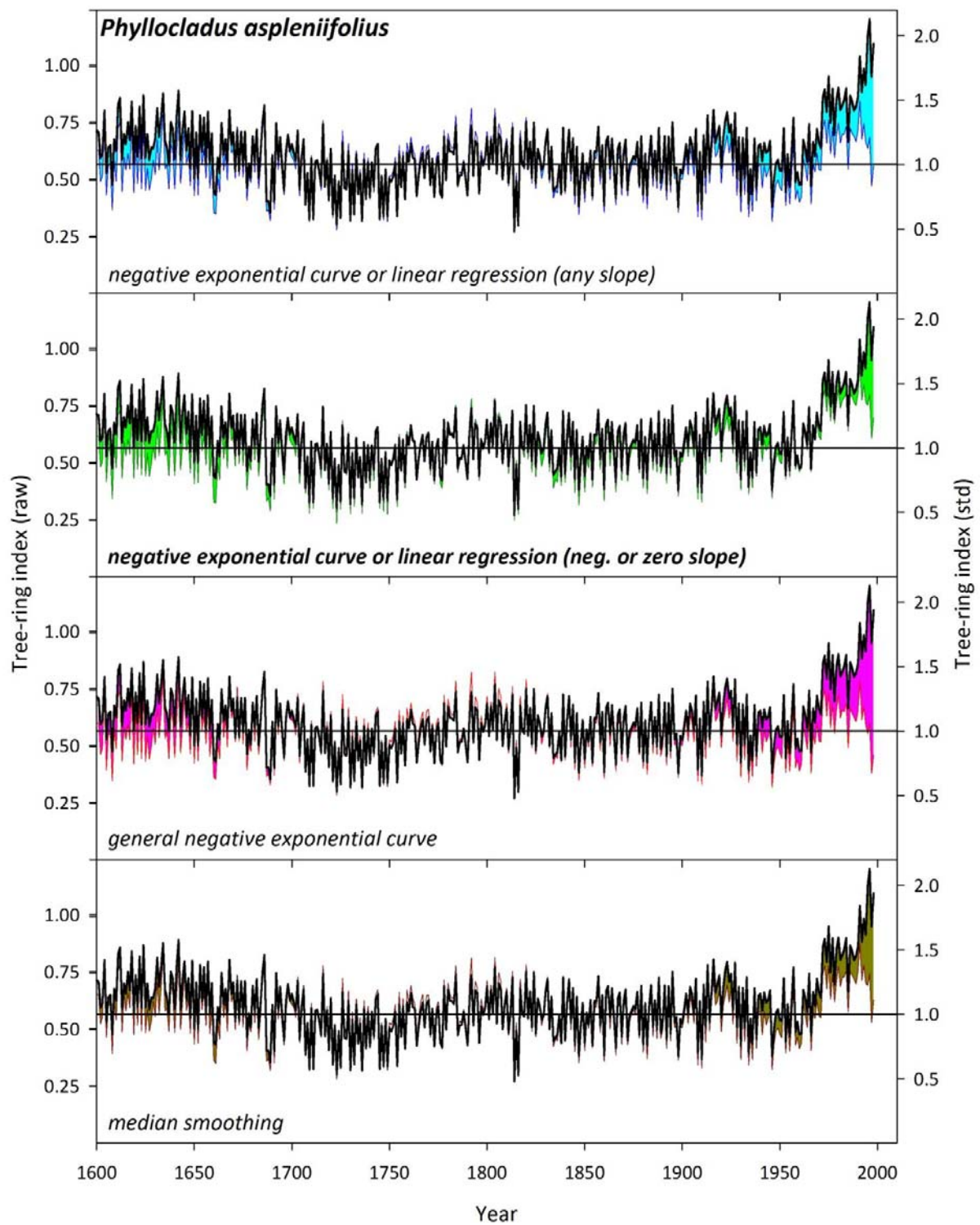
SI_Fig. 2. Same as in Fig. 1 for *Araucaria araucana*.



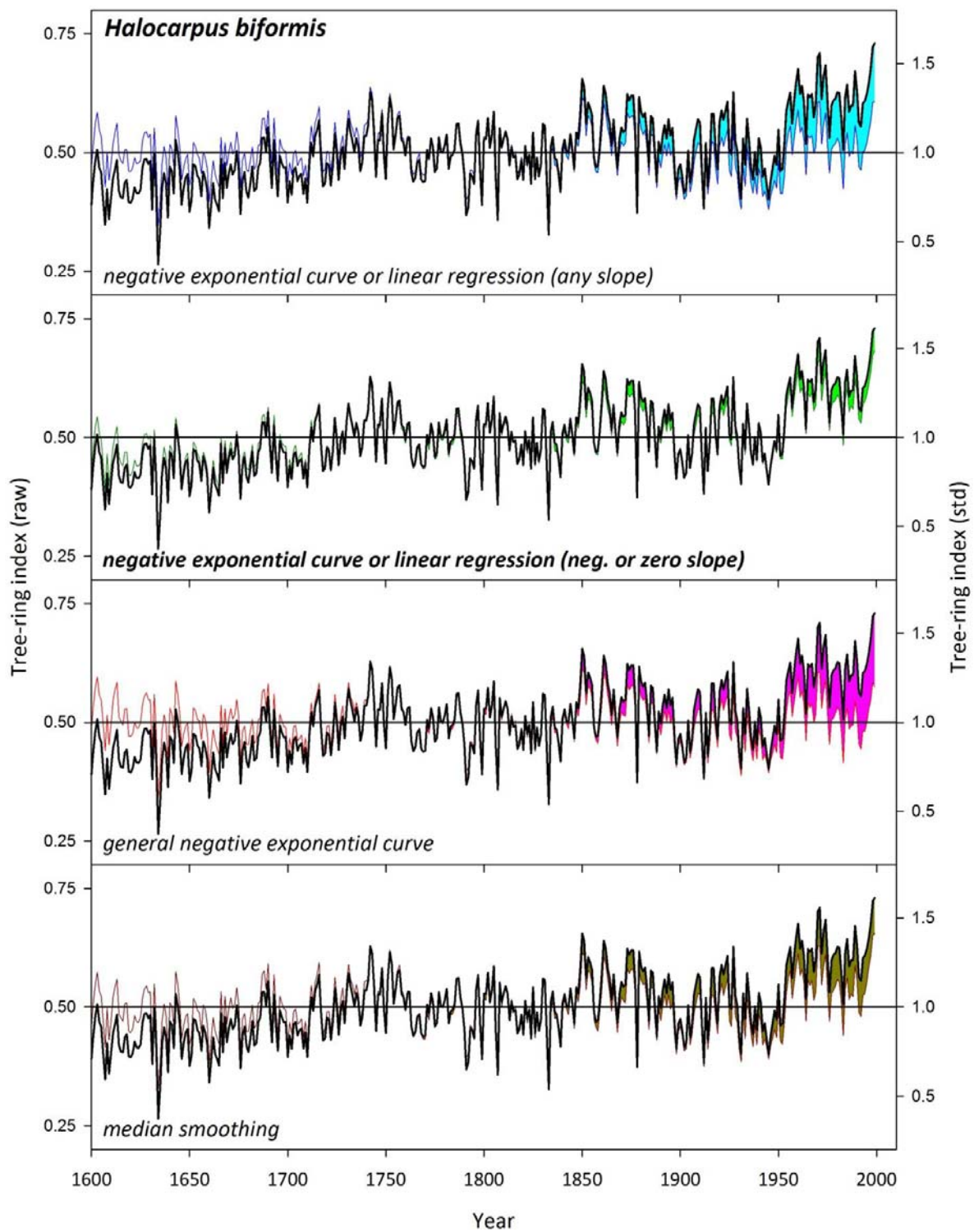
SI_Fig. 3. Same as in Fig. 1 for *Nothofagus betuloides*.



SI_Fig 4. Comparison between the *Lagarostrobos franklinii* raw tree-ring chronology without standardization (black line) and the chronologies produced using different curve fitting methods. The negative exponential curve or lineal regression with negative or zero slope (**highlighted**) captures most of the variance in the low-signal frequency since 1950 contained in the raw chronology.



SI_Fig 5. Same as in Fig. 4 for *Phyllocladus aspleniifolius*.



SI_Fig. 6. Same as in Fig. 4 for *Halocarpus biformis*.

Supplementary Information References

1. LaMarche, V.C., Holmes, R.L., Donwiddie, P. & Drew, L. Tree-ring chronologies of the Southern Hemisphere: 1. Argentina. Chronology Series V, vol 1. University of Arizona, Tucson. (1979).
2. Le Quesne, C., Acuña, C., Boninsegna, J.A., Rivera, A. & Barichivich, J. Long-term glacier variations in the Central Andes of Argentina and Chile, inferred from historical records and tree-ring reconstructed precipitation. *Paleogeogr. Paleoclim. Paleoecol.* **281**, 334-344 (2009).
3. Christie, D.A. Aridity changes in the Temperate-Mediterranean transition of the Andes since AD 1346 reconstructed from tree-rings. *Clim. Dyn.* **36**, 1505-1521 (2011).
4. Villalba, R. et al. Tree-ring based reconstructions of northern Patagonia precipitation since AD 1600. *The Holocene* **8**, 659-674 (1998).
5. Mundo, I.A., Roig Juñent, F.A., Villalba, R., Kitzberger, T. & Barrera, M.D. *Araucaria araucana* tree-ring chronologies in Argentina: spatial growth variations and climate influences. *Trees* DOI 10.1007/s00468-011-0605-3 (2011).
6. Villalba, R., Boninsegna, J.A. & Cobos, D.R. A tree-ring reconstruction of summer temperature between A.D. 1500 and 1974 in western Argentina. in: Third international conference on Southern Hemisphere Meteorology & Oceanography, Buenos Aires (1989).
7. Boninsegna, J.A., Keegan, J., Jacoby, G.C., D'Arrigo, R. & Holmes, R.L. Dendrochronological studies in Tierra del Fuego, Argentina. *Quat. South. Am.* **7**, 315-326 (1989).
8. Llancabure, J.J. Relaciones entre el crecimiento de *Nothofagus betuloides* y el clima local y de gran escala en bosques subantárticos de la Isla Navarino. Tesis Ingeniero Forestal, Facultad de Ciencias Forestales y Recursos Naturales, Universidad Austral de Chile (2011).
9. Cook, E.R. et al. Millennia-long tree-ring records from Tasmania and New Zealand: A basis for modelling climate variability and forcing, past, present and future. *J. Quaternary Sci.* **21**, 689-699 (2006).
10. Buckley, B.M., Cook, E.R., Peterson, M.J. & Barbetti, M. A changing temperature response with elevation for *Lagarostrobos franklinii* in Tasmania, Australia. *Clim. Change* **36**, 477-498 (1997).
11. Allen, K.J., Cook, E.R., Francey, R.J. & Michael, K. The climatic response of *Phyllocladus asplnifolius* (Labill.) Hook. f in Tasmania. *J. Biogeogr.* **28**, 305-316 (2001).
12. Allen, K.J. The temperature response in the ring width of *Phyllocladus asplnifolius* (Celery-top pine) along an altitudinal gradient in the Warra LTER area, Tasmania. *Austr. Geograp. Studies* **40**, 287-299 (2002).

13. D'Arrigo, R.D. et al. Temperature-sensitive tree-ring records from New Zealand: long-term context for recent warming trend. *Clim. Dyn.* **14**, 191-199 (1998).
14. Duncan, R.P., Fenwick, P., Palmer, J.P., McGlone, M.S., & Turney, C. Non-uniform interhemispheric temperature trends over the past 550 years. *Clim. Dyn.* **35**, 1429-1438 (2010).
15. Briffa, K.R. in *Analysis of Climate Variability: Applications of Statistical Techniques*. (eds. Von Storch, H. & Navarra, A.) 77-94 (Springer, Berlin, 1995).
16. Wigley, T.M.L., Briffa, K.R. & Jones, P.D. On the average value of correlated time series with applications in dendroclimatology and hydrometeorology. *J. Clim. Appl. Meteorol.* **23**, 201-213 (1984).
17. Trenberth, K.E. Signal versus Noise in the Southern Oscillation. *Mon. Weather. Rev.* **112**:326-332 (1984).
18. Cook, E.R. A time series analysis approach to tree ring standardization. PhD. Dissertation, University of Arizona Tucson, 185 pp. (1985).
19. Marshall, G. J. Trends in the southern annular mode from observations and reanalyses. *J. Climate* **16**, 4134-4143 (2003).
20. Thompson, D. W. J. & Wallace, J. M. Annular modes in the extratropical circulation. Part I: Month-to-month variability. *J. Climate* **13**, 1000-1016 (2000).
21. Fogt, R. L. et al. Historical SAM variability. Part II: Twentieth century variability and trends from reconstructions, observations, and the IPCC AR4 models. *J. Climate*, **22**, 5346-5365 (2009).
22. Jones, J. M. et al. Historical SAM variability. Part I: Century-length seasonal reconstructions. *J. Climate* **22**, 5319-5345 (2009).